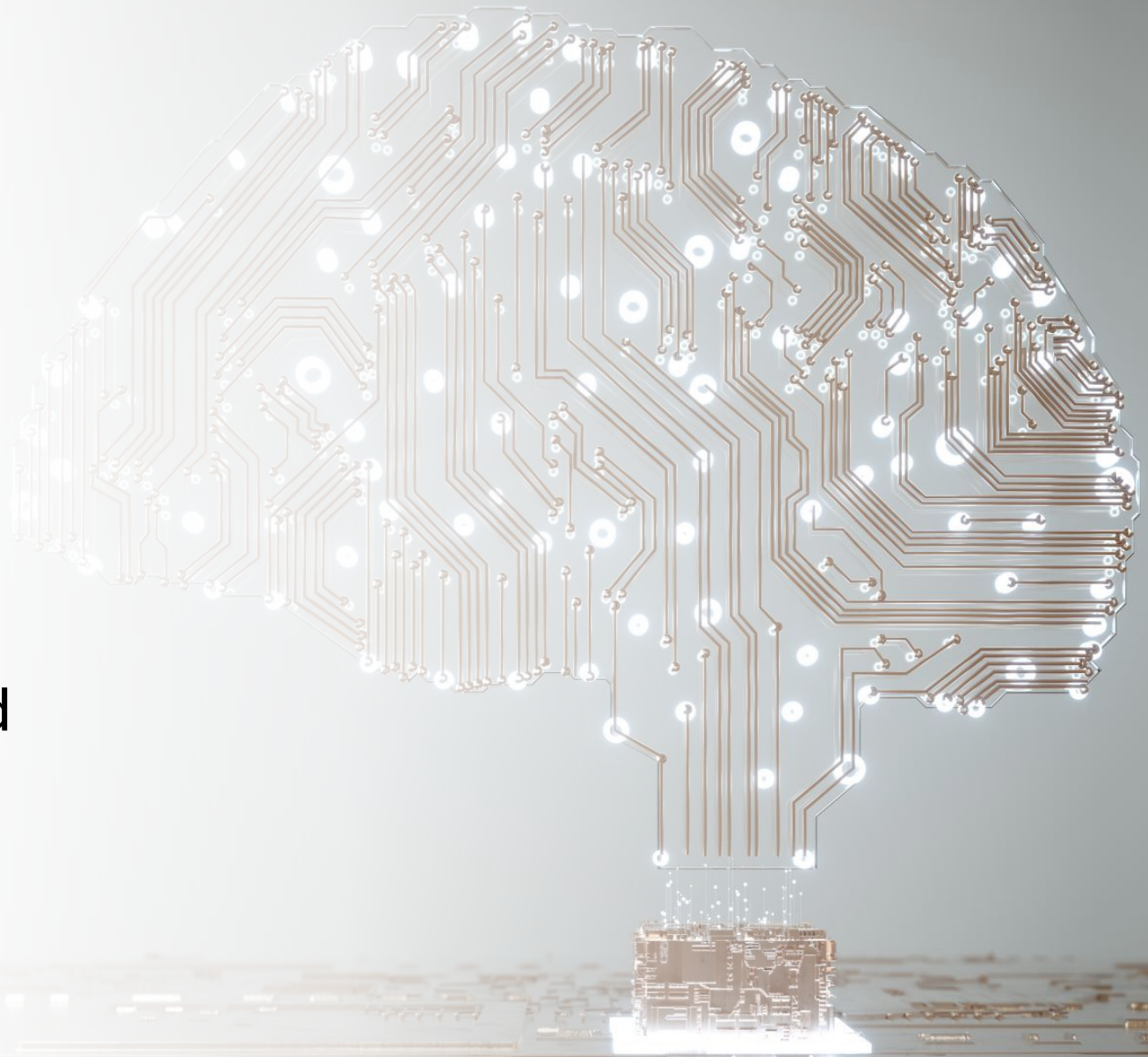


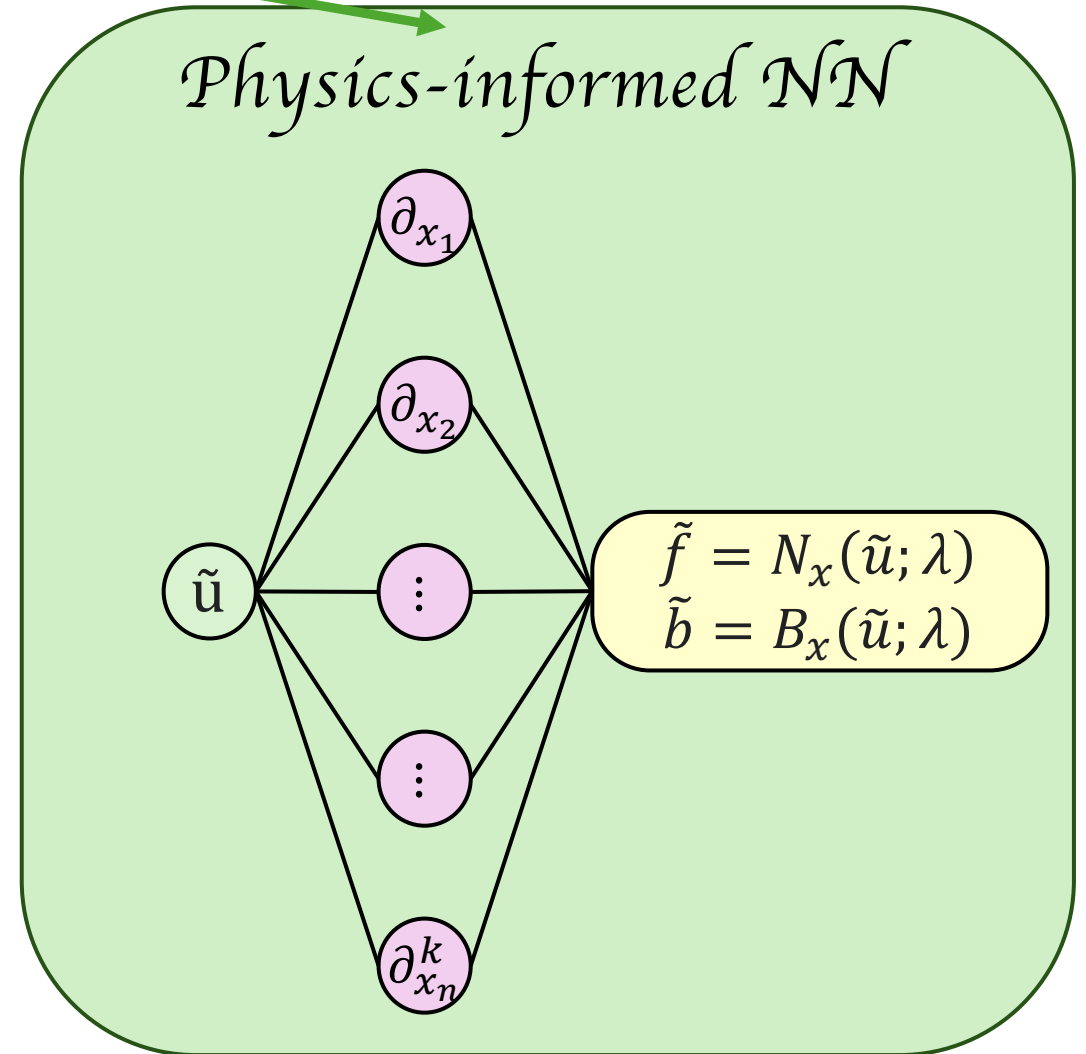
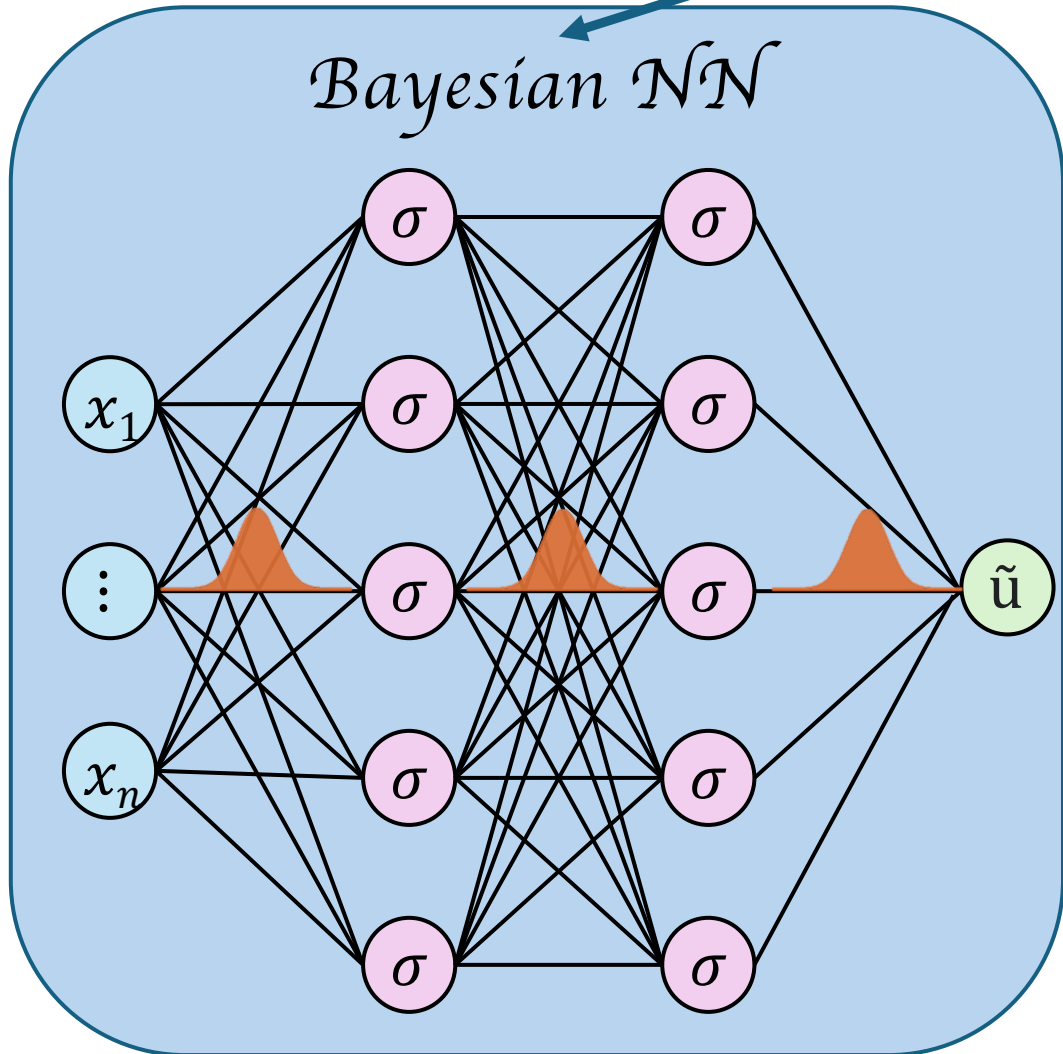
B-PINNs

"to solve linear and nonlinear PDEs
with noisy data for both forward and
inverse problems"

Paper by: Liu Yang, Xuhui Meng, George Em Karniadakis
Presented by: Sonja Joost, Gohar Tamrazyan



B-PINN



Plan



Motivation



Framework



Problem Approach



Sampling Approaches



Results

The Problem - Form of a general PDE



$$\mathcal{N}_x(\mathbf{u}; \lambda) = f, \quad \mathbf{x} \in D$$

$$\mathcal{B}_x(\mathbf{u}; \lambda) = b, \quad \mathbf{x} \in \Gamma$$

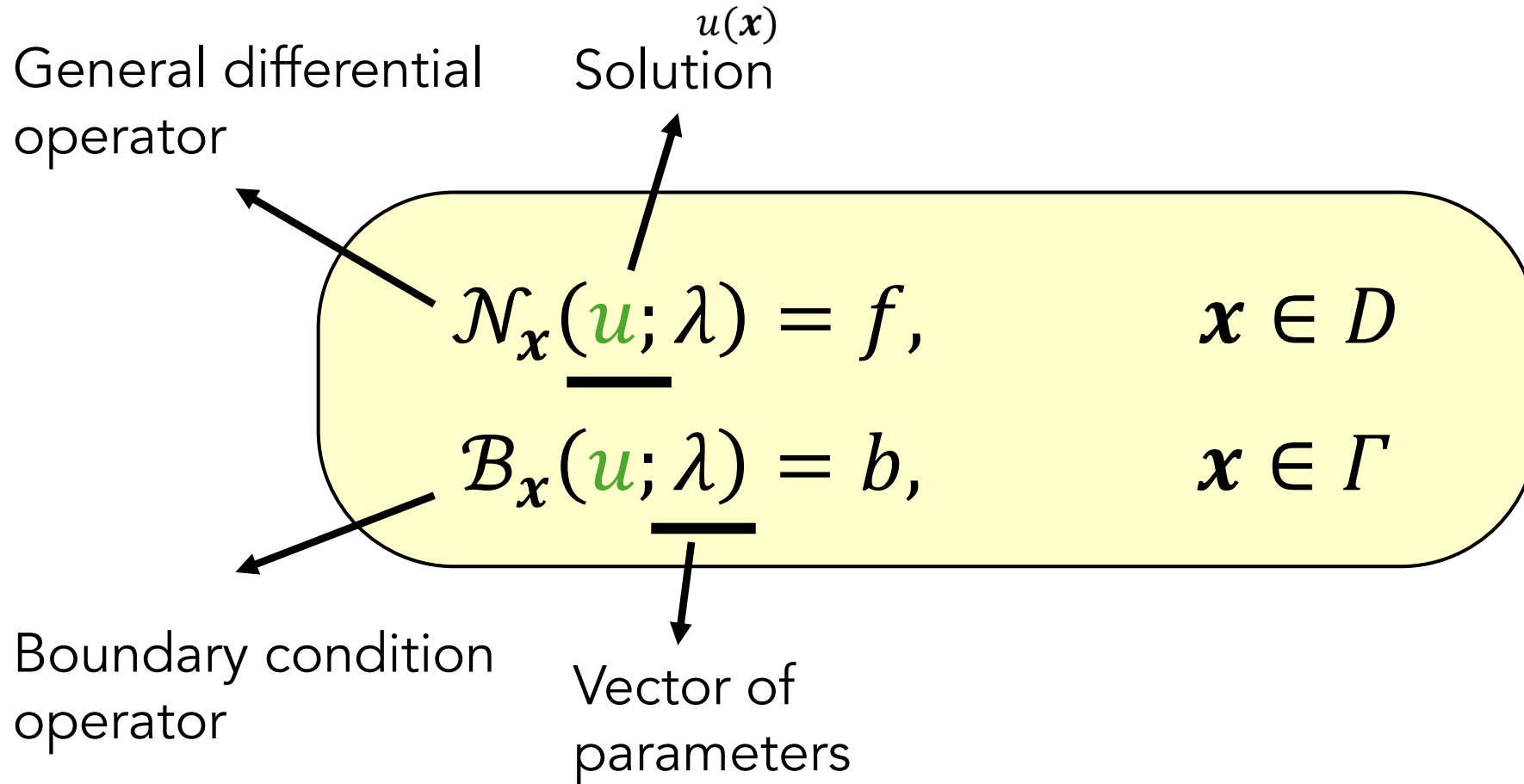
The Problem - Form of a general PDE

General differential operator

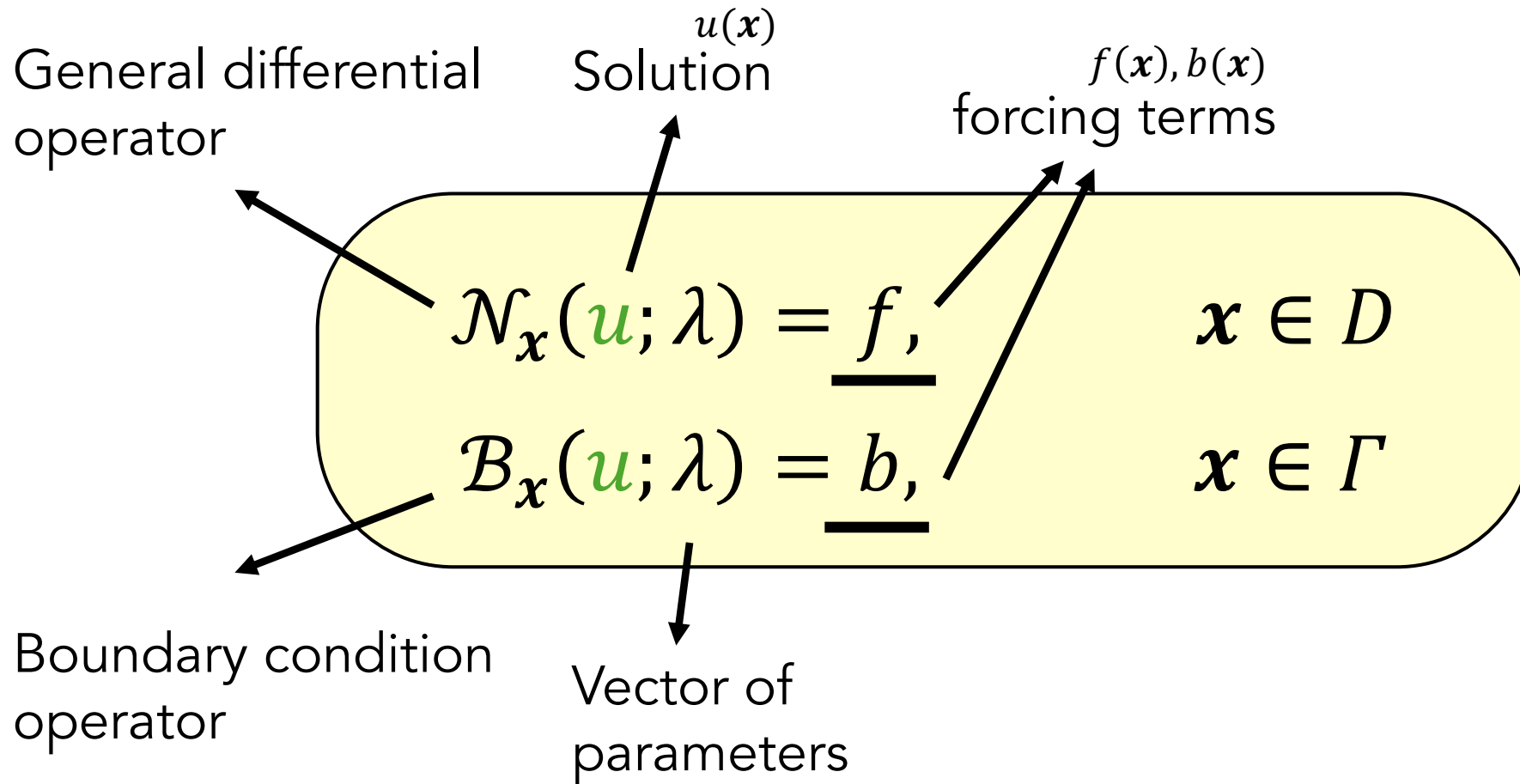
$$\begin{aligned} \underline{\mathcal{N}_x}(\mathbf{u}; \lambda) &= f, & \mathbf{x} \in D \\ \underline{\mathcal{B}_x}(\mathbf{u}; \lambda) &= b, & \mathbf{x} \in \Gamma \end{aligned}$$

Boundary condition operator

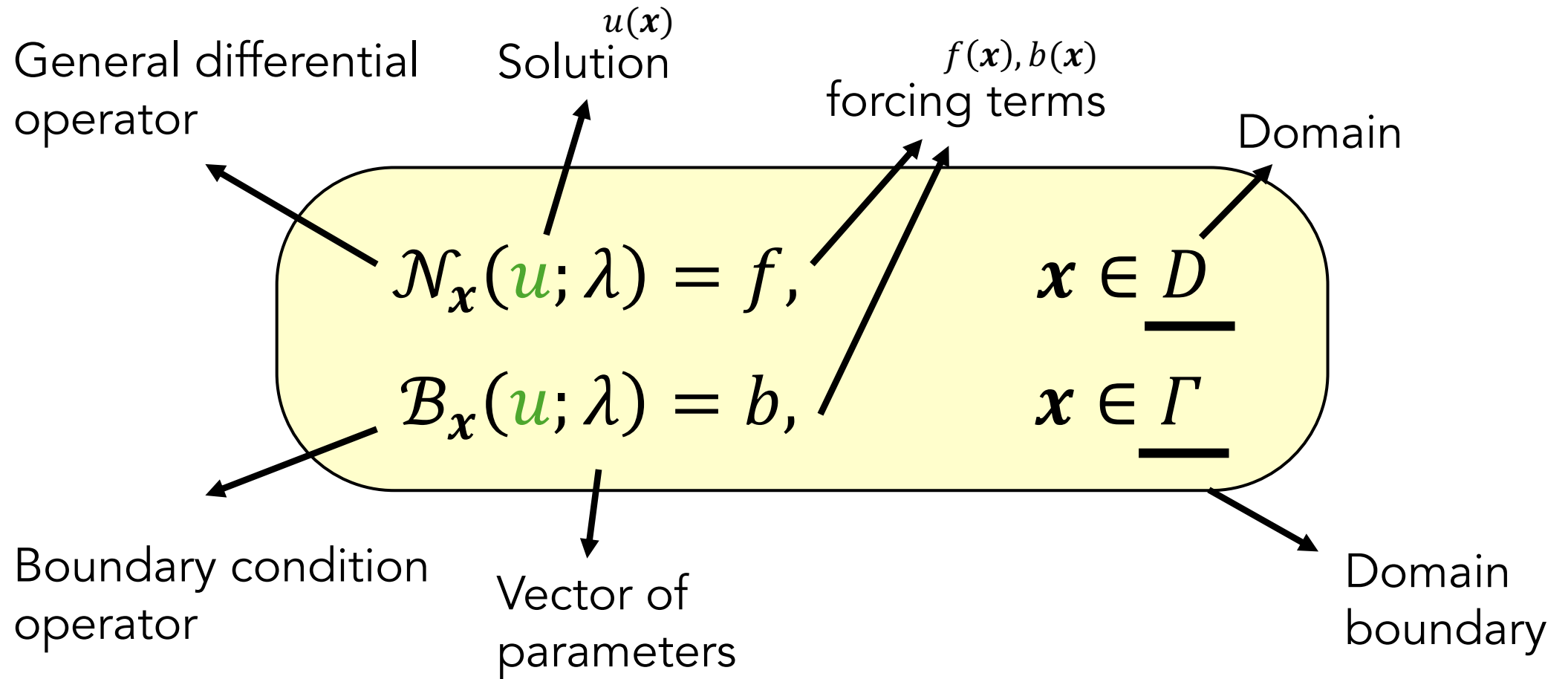
The Problem - Form of a general PDE



The Problem - Form of a general PDE



The Problem - Form of a general PDE





Framework

What we have and what we are looking for.

Dataset \mathcal{D}

$$\mathcal{D} = \mathcal{D}_u \cup \mathcal{D}_f \cup \mathcal{D}_b$$

$$\mathcal{D}_u = \left\{ \left(\mathbf{x}_u^{(i)}, \underline{\bar{u}}^{(i)} \right) \right\}_{i=1}^{N_u}$$

$$\mathcal{D}_f = \left\{ \left(\mathbf{x}_f^{(i)}, \underline{\bar{f}}^{(i)} \right) \right\}_{i=1}^{N_f}$$

$$\mathcal{D}_b = \left\{ \left(\mathbf{x}_b^{(i)}, \underline{\bar{b}}^{(i)} \right) \right\}_{i=1}^{N_b}$$

$$\mathcal{N}_x(\mathbf{u}; \lambda) = f, \quad \mathbf{x} \in D$$

$$\mathcal{B}_x(\mathbf{u}; \lambda) = b, \quad \mathbf{x} \in \Gamma$$

Dataset \mathcal{D}

$$\begin{aligned} \mathcal{N}_x(\mathbf{u}; \lambda) &= f, & x \in D \\ \mathcal{B}_x(\mathbf{u}; \lambda) &= b, & x \in \Gamma \end{aligned}$$

Measurements

$$\mathcal{D} = \mathcal{D}_u \cup \mathcal{D}_f \cup \mathcal{D}_b$$

$$\mathcal{D}_u = \left\{ \left(\mathbf{x}_u^{(i)}, \underline{\bar{u}}^{(i)} \right) \right\}_{i=1}^{N_u}$$

$$\mathcal{D}_f = \left\{ \left(\mathbf{x}_f^{(i)}, \underline{\bar{f}}^{(i)} \right) \right\}_{i=1}^{N_f}$$

$$\mathcal{D}_b = \left\{ \left(\mathbf{x}_b^{(i)}, \underline{\bar{b}}^{(i)} \right) \right\}_{i=1}^{N_b}$$

$$\bar{u}^{(i)} = u(\mathbf{x}_u^{(i)}) + \epsilon_u^{(i)}, \quad i = 1, 2, \dots, N_u$$

$$\bar{f}^{(i)} = f(\mathbf{x}_f^{(i)}) + \epsilon_f^{(i)}, \quad i = 1, 2, \dots, N_f$$

$$\bar{b}^{(i)} = b(\mathbf{x}_b^{(i)}) + \epsilon_b^{(i)}, \quad i = 1, 2, \dots, N_b$$

Dataset \mathcal{D}

$$\begin{aligned} \mathcal{N}_x(\mathbf{u}; \lambda) &= f, & x \in D \\ \mathcal{B}_x(\mathbf{u}; \lambda) &= b, & x \in \Gamma \end{aligned}$$

$$\mathcal{D} = \mathcal{D}_u \cup \mathcal{D}_f \cup \mathcal{D}_b$$

$$\mathcal{D}_u = \left\{ \left(\mathbf{x}_u^{(i)}, \bar{u}^{(i)} \right) \right\}_{i=1}^{N_u}$$

$$\mathcal{D}_f = \left\{ \left(\mathbf{x}_f^{(i)}, \bar{f}^{(i)} \right) \right\}_{i=1}^{N_f}$$

$$\mathcal{D}_b = \left\{ \left(\mathbf{x}_b^{(i)}, \bar{b}^{(i)} \right) \right\}_{i=1}^{N_b}$$

Measurements

Hidden real Values

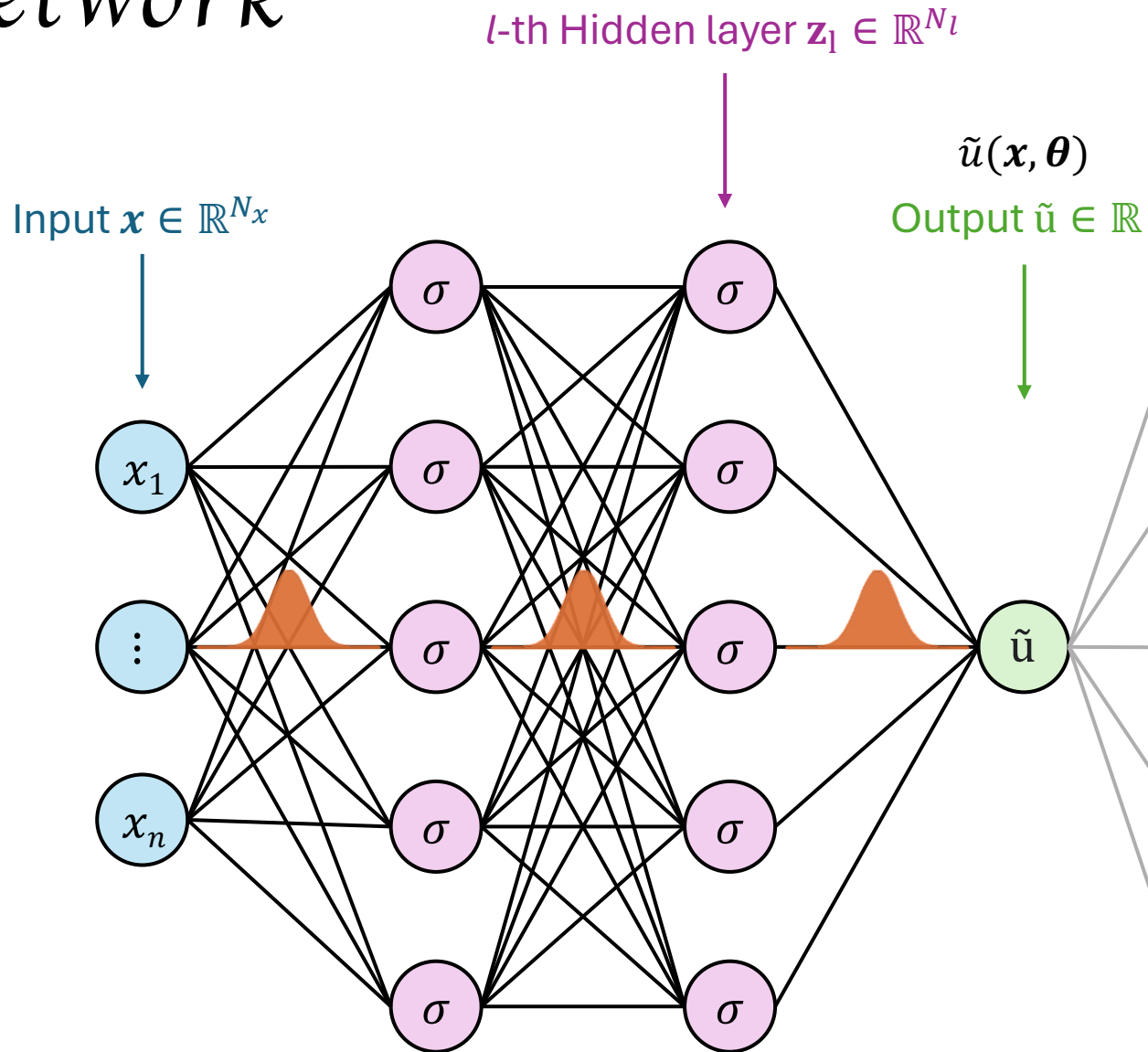
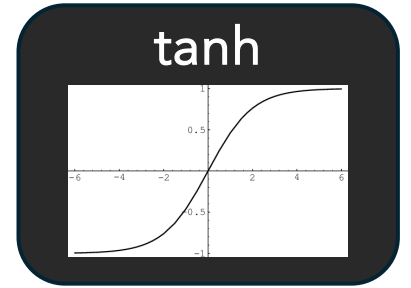
Independent Gaussian noises with zero mean

$$\bar{u}^{(i)} = u(\mathbf{x}_u^{(i)}) + \epsilon_u^{(i)}, \quad i = 1, 2, \dots, N_u$$

$$\bar{f}^{(i)} = f(\mathbf{x}_f^{(i)}) + \epsilon_f^{(i)}, \quad i = 1, 2, \dots, N_f$$

$$\bar{b}^{(i)} = b(\mathbf{x}_b^{(i)}) + \epsilon_b^{(i)}, \quad i = 1, 2, \dots, N_b$$

Network



$$\begin{aligned} \tilde{f} &= N_x(\tilde{u}; \lambda) \\ \tilde{b} &= B_x(\tilde{u}; \lambda) \end{aligned}$$



Problem Approach

Problem approach

1. Assume Prior distribution

$$P(\boldsymbol{\theta})$$

2. Compute Likelihood

$$P(\mathcal{D} | \boldsymbol{\theta})$$

3. Compute Posterior Distribution

$$P(\boldsymbol{\theta} | \mathcal{D})$$



Problem approach

1. Assume Prior distribution

$$P(\boldsymbol{\theta})$$

2. Compute Likelihood

$$P(\mathcal{D} \mid \boldsymbol{\theta})$$

3. Compute Posterior Distribution

$$P(\boldsymbol{\theta} \mid \mathcal{D})$$

4. Sample from the Posterior Distribution

$$\{\boldsymbol{\theta}^{(i)}\}_{i=1}^M$$

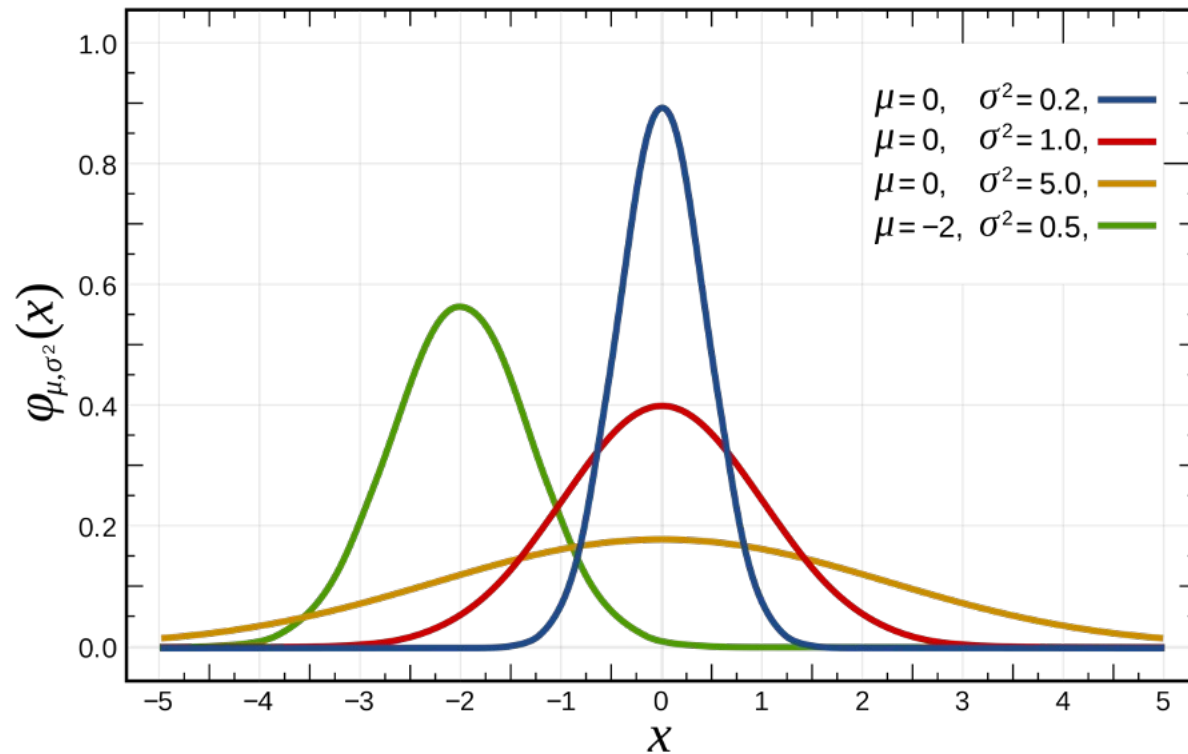
5. Obtain statistics from samples

$$\{\tilde{u}(\mathbf{x}, \boldsymbol{\theta}^{(i)})\}_{i=1}^M$$



1. Prior Distribution

Gaussian Distributions with zero mean



2. Likelihood

$$P(\mathcal{D} \mid \boldsymbol{\theta}) = P(D_u \cup D_f \cup D_b \mid \boldsymbol{\theta}) = P(\mathcal{D}_u \mid \boldsymbol{\theta})P(\mathcal{D}_f \mid \boldsymbol{\theta})P(\mathcal{D}_b \mid \boldsymbol{\theta})$$



2. Likelihood

$$P(\mathcal{D} \mid \boldsymbol{\theta}) = P(D_u \cup D_f \cup D_b \mid \boldsymbol{\theta}) = P(D_u \mid \boldsymbol{\theta})P(D_f \mid \boldsymbol{\theta})P(D_b \mid \boldsymbol{\theta})$$

$$P(D_u \mid \boldsymbol{\theta}) = P\left(\left\{\left(\mathbf{x}_u^{(i)}, \bar{u}^{(i)}\right)\right\}_{i=1}^{N_u} \mid \boldsymbol{\theta}\right) = \prod_{i=1}^{N_u} \frac{1}{\sqrt{2\pi\sigma_u^{(i)2}}} \exp\left(-\frac{\left(\tilde{u}\left(\mathbf{x}_u^{(i)}; \boldsymbol{\theta}\right) - \bar{u}^{(i)}\right)^2}{2\sigma_u^{(i)2}}\right)$$

2. Likelihood

Prior of θ is a Gaussian Distribution

$$P(\mathcal{D} | \theta) = P(D_u \cup D_f \cup D_b | \theta) = P(D_u | \theta)P(D_f | \theta)P(D_b | \theta)$$

$$P(D_u | \theta) = P\left(\left\{\left(x_u^{(i)}, \bar{u}^{(i)}\right)\right\}_{i=1}^{N_u} | \theta\right) = \prod_{i=1}^{N_u} \frac{1}{\sqrt{2\pi\sigma_u^{(i)2}} \exp\left(-\frac{\left(\tilde{u}\left(x_u^{(i)}; \theta\right) - \bar{u}^{(i)}\right)^2}{2\sigma_u^{(i)2}}\right)}$$

2. Likelihood

Computationally heavy

Prior of θ is a Gaussian Distribution

$$P(\mathcal{D} | \theta) = P(D_u \cup D_f \cup D_b | \theta) = P(\mathcal{D}_u | \theta)P(\mathcal{D}_f | \theta)P(\mathcal{D}_b | \theta)$$

$$P(\mathcal{D}_u | \theta) = P\left(\left\{\left(x_u^{(i)}, \bar{u}^{(i)}\right)\right\}_{i=1}^{N_u} | \theta\right) = \prod_{i=1}^{N_u} \frac{1}{\sqrt{2\pi\sigma_u^{(i)2}}} \exp\left(-\frac{\left(\tilde{u}\left(x_u^{(i)}; \theta\right) - \bar{u}^{(i)}\right)^2}{2\sigma_u^{(i)2}}\right)$$

$$P(\mathcal{D}_f | \theta) = P\left(\left\{\left(x_f^{(i)}, \bar{f}^{(i)}\right)\right\}_{i=1}^{N_f} | \theta\right) = \prod_{i=1}^{N_f} \frac{1}{\sqrt{2\pi\sigma_f^{(i)2}}} \exp\left(-\frac{\left(\tilde{f}\left(x_f^{(i)}; \theta\right) - \bar{f}^{(i)}\right)^2}{2\sigma_f^{(i)2}}\right)$$

$$P(\mathcal{D}_b | \theta) = P\left(\left\{\left(x_b^{(i)}, \bar{b}^{(i)}\right)\right\}_{i=1}^{N_b} | \theta\right) = \prod_{i=1}^{N_b} \frac{1}{\sqrt{2\pi\sigma_b^{(i)2}}} \exp\left(-\frac{\left(\tilde{b}\left(x_b^{(i)}; \theta\right) - \bar{b}^{(i)}\right)^2}{2\sigma_b^{(i)2}}\right)$$

3. Posterior Distribution

$$P(\theta|\mathcal{D}) = \frac{P(\mathcal{D}|\theta)P(\theta)}{P(\mathcal{D})} \simeq P(\mathcal{D}|\theta)P(\theta) = \textit{likelihood} \times \textit{prior}$$

Bayes' Theorem

Equal up to a constant





Sampling Approaches

HMC and VI

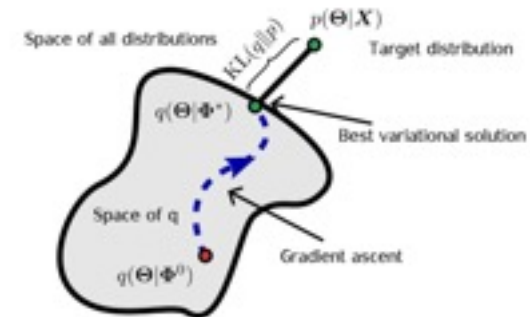
4. Posterior Sampling Approaches

HMC - Hamiltonian Monte Carlo

- Complex probability distributions
- High-dimensional parameter spaces
- Hamiltonian dynamics
- Parameters of interest – positions θ
- Auxiliary momentum variable r

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VI - Variational Inference



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HMC - Hamiltonian Monte Carlo

- Complex probability distributions
- High-dimensional parameter spaces
- Hamiltonian dynamics

- Parameters of interest – positions θ
- Auxiliary momentum variable r

HMC - Setup

$$P(\boldsymbol{\theta}|\mathcal{D}) \simeq P(\mathcal{D}|\boldsymbol{\theta}) \cdot P(\boldsymbol{\theta})$$

$$= \exp(\ln(P(\mathcal{D}|\boldsymbol{\theta}) \cdot P(\boldsymbol{\theta})))$$

$$= \exp(\ln(P(\mathcal{D}|\boldsymbol{\theta})) + \ln(P(\boldsymbol{\theta})))$$

$$= \exp\left(-\left(-\ln(P(\mathcal{D}|\boldsymbol{\theta})) - \ln(P(\boldsymbol{\theta}))\right)\right)$$

$$= \exp(-U(\boldsymbol{\theta}))$$

HMC - Setup

$$U(\boldsymbol{\theta}) = (-\ln(P(\mathcal{D}|\boldsymbol{\theta})) - \ln(P(\boldsymbol{\theta})))$$

$$P(\boldsymbol{\theta}|\mathcal{D}) \simeq \exp(-U(\boldsymbol{\theta}))$$

joint distribution $\boldsymbol{\pi}$:

$$\boldsymbol{\pi}(\boldsymbol{\theta}, \mathbf{r}) \sim \exp\left(-U(\boldsymbol{\theta}) - \frac{1}{2}\mathbf{r}^T \mathbf{M}^{-1}\mathbf{r}\right)$$

Hamiltonian system:

$$H(\boldsymbol{\theta}, \mathbf{r}) = U(\boldsymbol{\theta}) + \frac{1}{2}\mathbf{r}^T \mathbf{M}^{-1}\mathbf{r}$$

Hamiltonian dynamics:

$$\frac{d\mathbf{r}}{dt} = -\frac{\partial H}{\partial \boldsymbol{\theta}}$$

$$\frac{d\boldsymbol{\theta}}{dt} = \frac{\partial H}{\partial \mathbf{r}}$$

$$d\mathbf{r} = -\nabla U(\boldsymbol{\theta})dt$$

$$d\boldsymbol{\theta} = \mathbf{M}^{-1}\mathbf{r}dt$$

HMC - Setup

Hamiltonian dynamics:

$$\frac{d\mathbf{r}}{dt} = -\frac{\partial H}{\partial \boldsymbol{\theta}}$$

$$d\mathbf{r} = -\nabla U(\boldsymbol{\theta})dt$$

$$\frac{d\boldsymbol{\theta}}{dt} = \frac{\partial H}{\partial \mathbf{r}}$$

$$d\boldsymbol{\theta} = \mathbf{M}^{-1}\mathbf{r}dt$$

$$\begin{aligned} & \frac{d}{dt}H(\boldsymbol{\theta}(t), \mathbf{r}(t)) \\ &= \frac{\partial H}{\partial \boldsymbol{\theta}} \cdot \frac{d\boldsymbol{\theta}}{dt} + \frac{\partial H}{\partial \mathbf{r}} \cdot \frac{d\mathbf{r}}{dt} \\ &= -\frac{d\mathbf{r}}{dt} \cdot \frac{d\boldsymbol{\theta}}{dt} + \frac{d\boldsymbol{\theta}}{dt} \cdot \frac{d\mathbf{r}}{dt} \\ &= 0 \end{aligned}$$

HMC

Leapfrog integration

Metropolis Hastings

Algorithm 1 Hamiltonian Monte Carlo.

Require: initial states for θ^{t_0} and time step size δt .

for $k = 1, 2 \dots N$ **do**

Sample $\mathbf{r}^{t_{k-1}}$ from $\mathcal{N}(0, \mathbf{M})$,

$(\theta_0, \mathbf{r}_0) \leftarrow (\theta^{t_{k-1}}, \mathbf{r}^{t_{k-1}})$.

for $i = 0, 1 \dots (L - 1)$ **do**

$\mathbf{r}_i \leftarrow \mathbf{r}_i - \frac{\delta t}{2} \nabla U(\theta_i)$,

$\theta_{i+1} \leftarrow \theta_i + \delta t \mathbf{M}^{-1} \mathbf{r}_i$,

$\mathbf{r}_{i+1} \leftarrow \mathbf{r}_i - \frac{\delta t}{2} \nabla U(\theta_{i+1})$,

end for

Metropolis-Hastings step:

Sample p from Uniform[0, 1],

$\alpha \leftarrow \min\{1, \exp(H(\theta_L, \mathbf{r}_L) - H(\theta^{t_{k-1}}, \mathbf{r}^{t_{k-1}}))\}$.

if $p \geq \alpha$ **then**

$\theta^{t_k} \leftarrow \theta_L$,

else

$\theta^{t_k} \leftarrow \theta^{t_{k-1}}$.

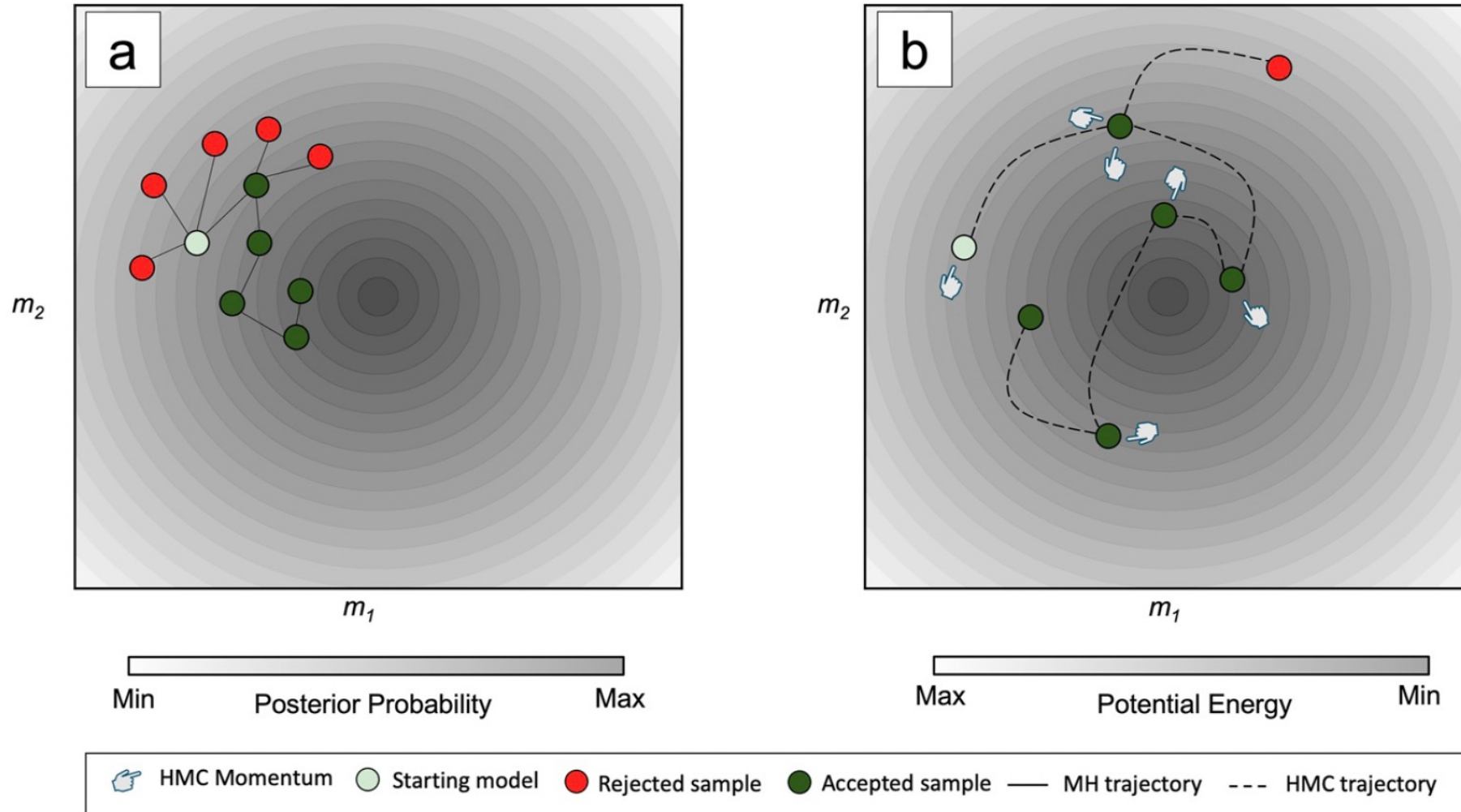
end if

end for

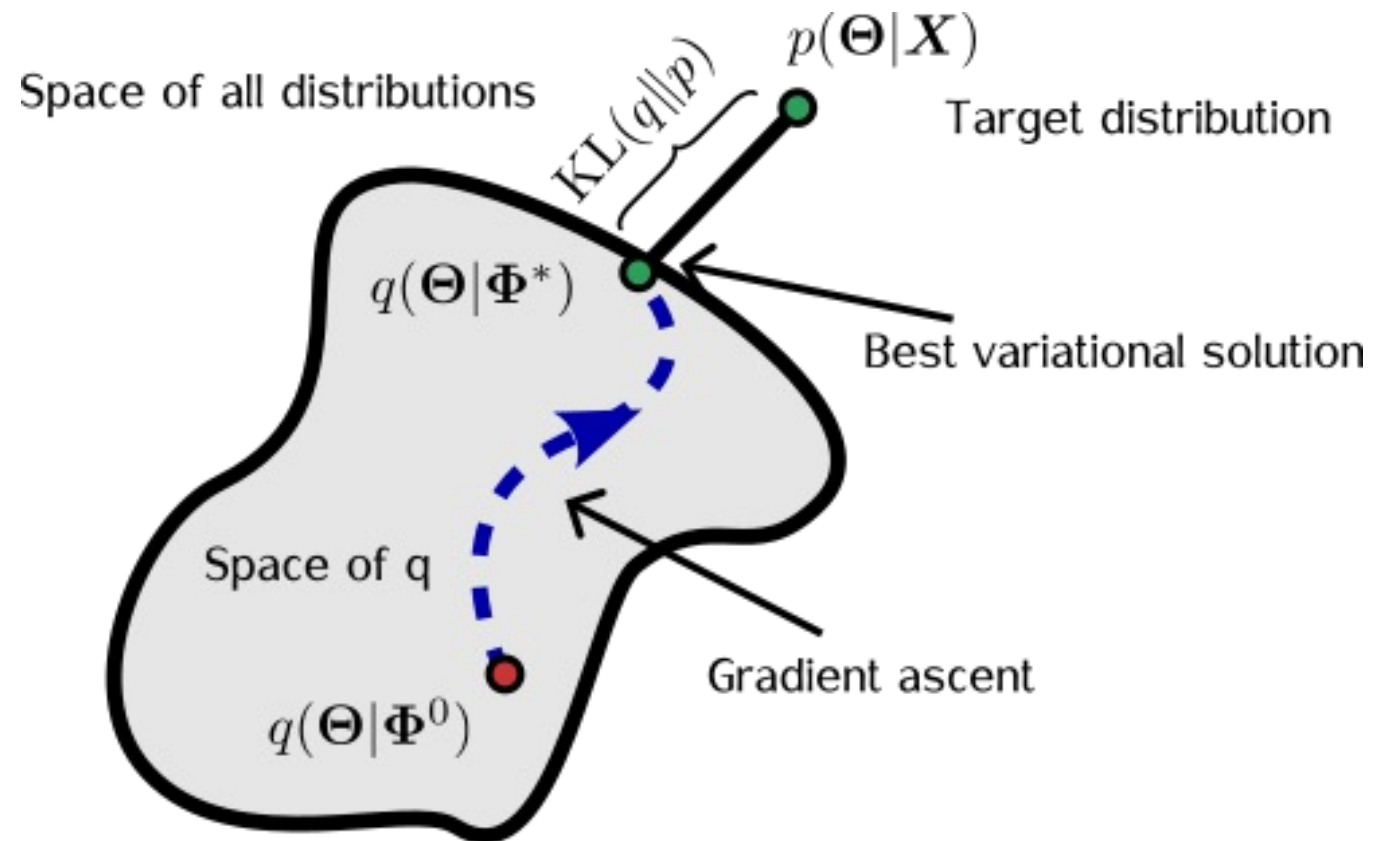
Calculate $\{\tilde{u}(\mathbf{x}, \theta^{t_{N+1-j}})\}_{j=1}^M$ as samples of $u(\mathbf{x})$, similarly for other terms.

HMC

$$H(\theta, r) = U(\theta) + \frac{1}{2} r^T M^{-1} r$$



VI - Variational Inference





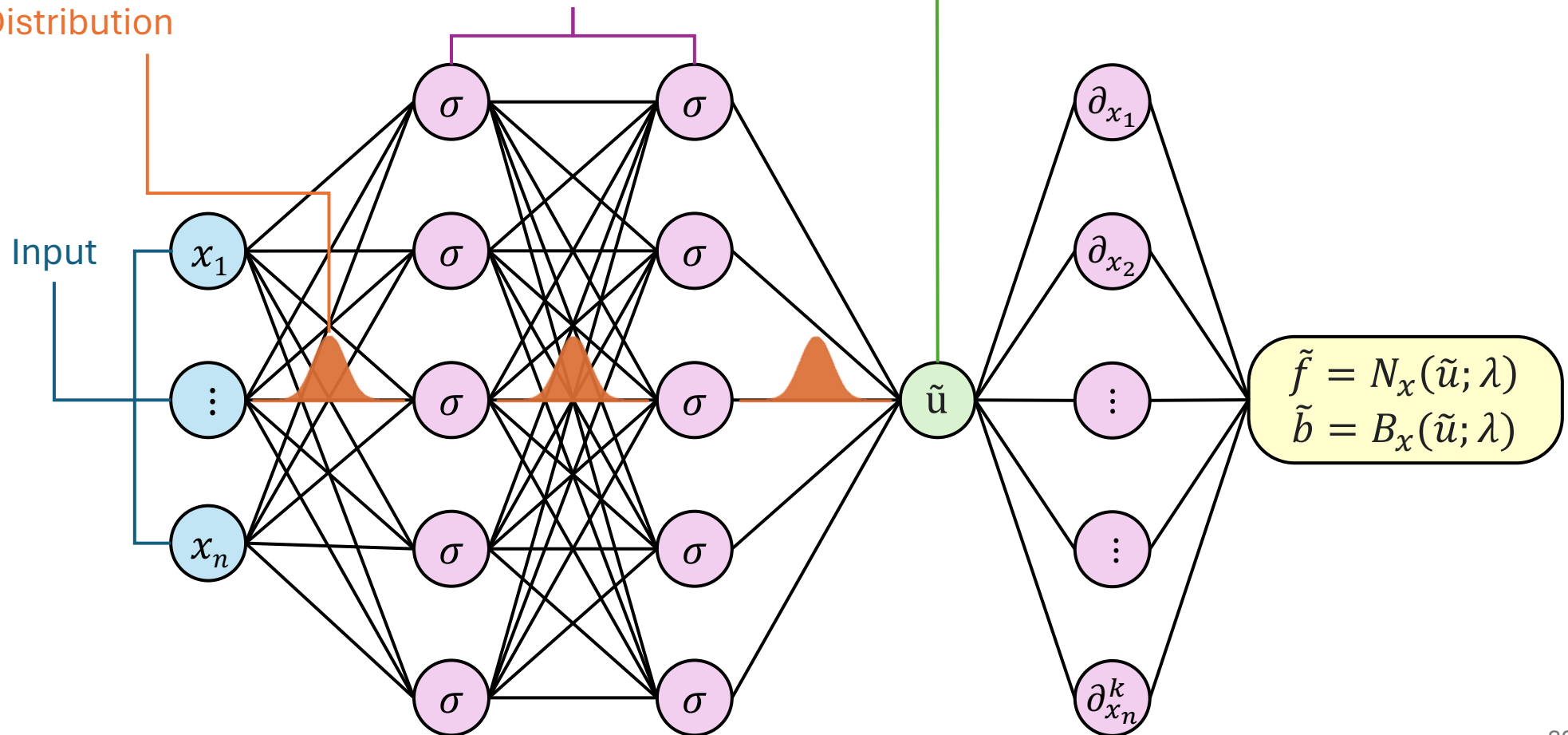
Results

The Network

Prior: Independent Standard Gaussian Distribution

2 hidden layers of width 50

Output



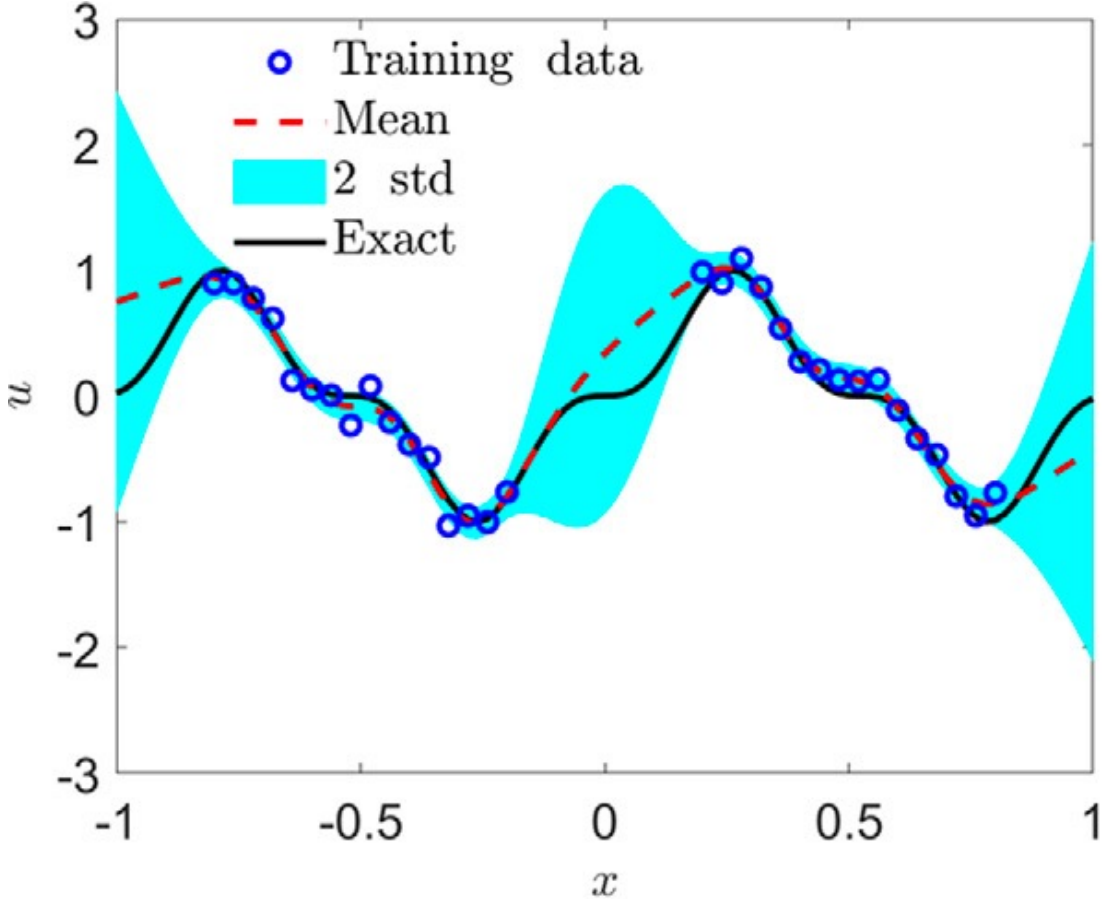
Function regression

$$u(x) = \sin^3(6x), \quad x \in [-1, 1]$$

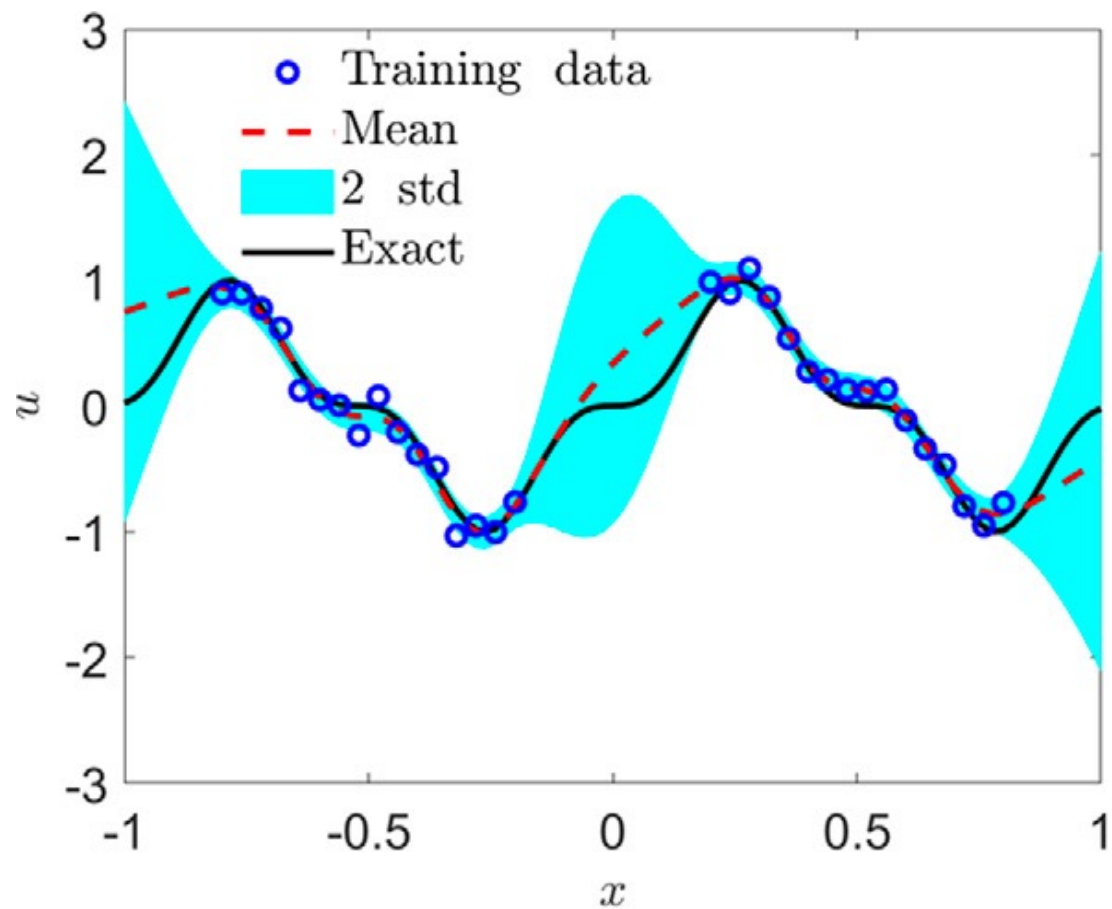
- BNN instead of B-PINN
- 32 training points $\bar{u}^{(i)}$ in $[-0.8, -0.2] \cup [0.2, 0.8]$



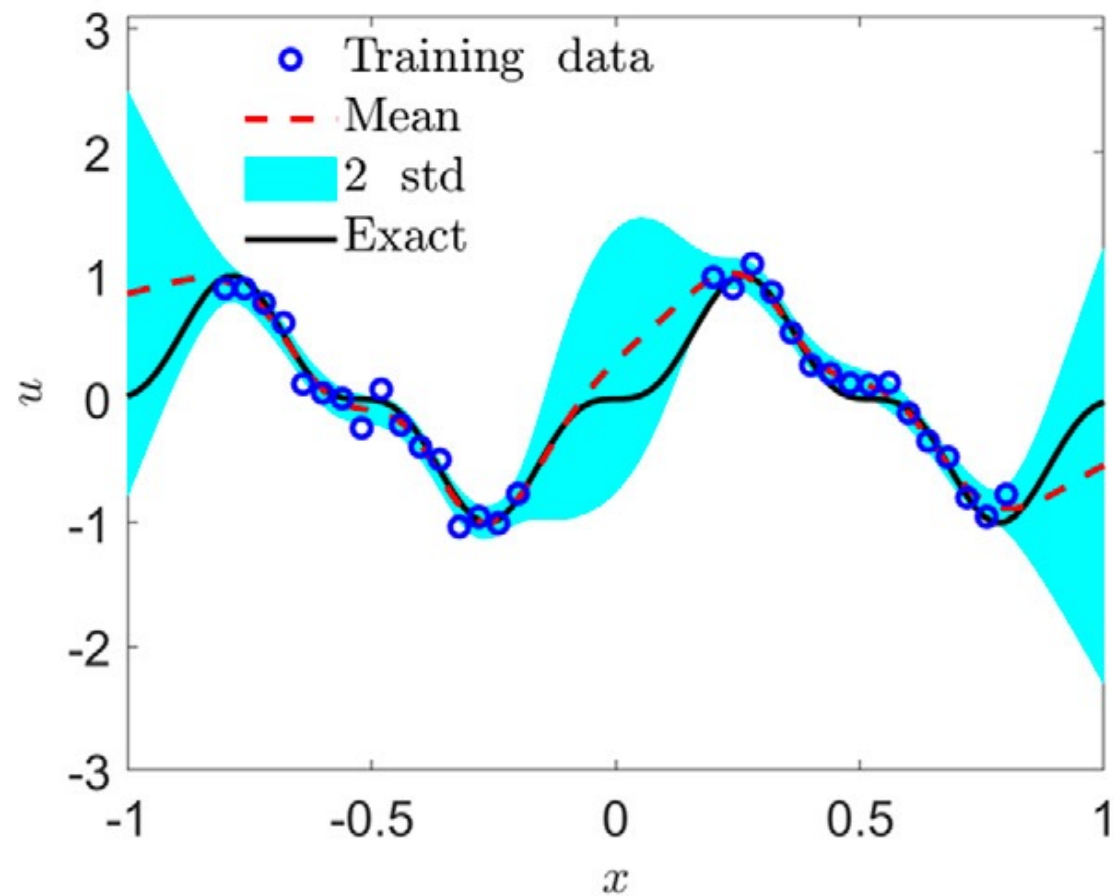
BNN-GPR



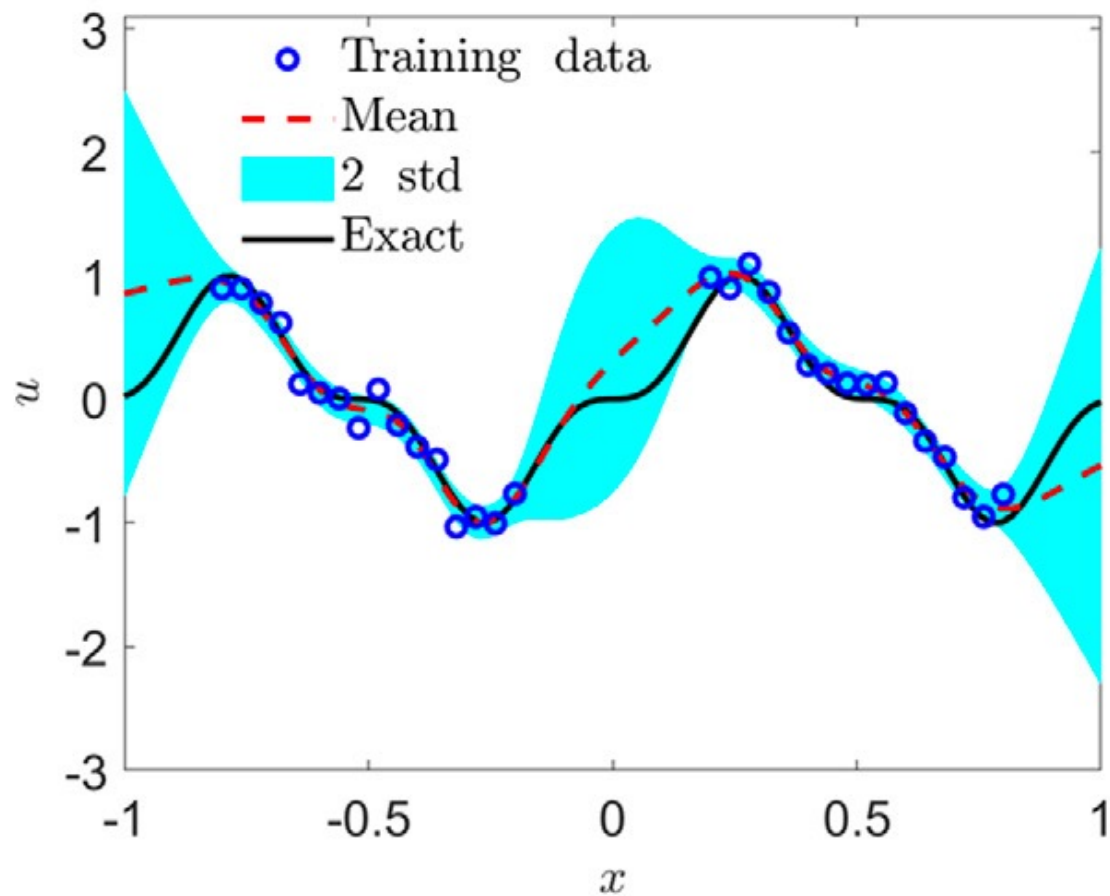
BNN-GPR



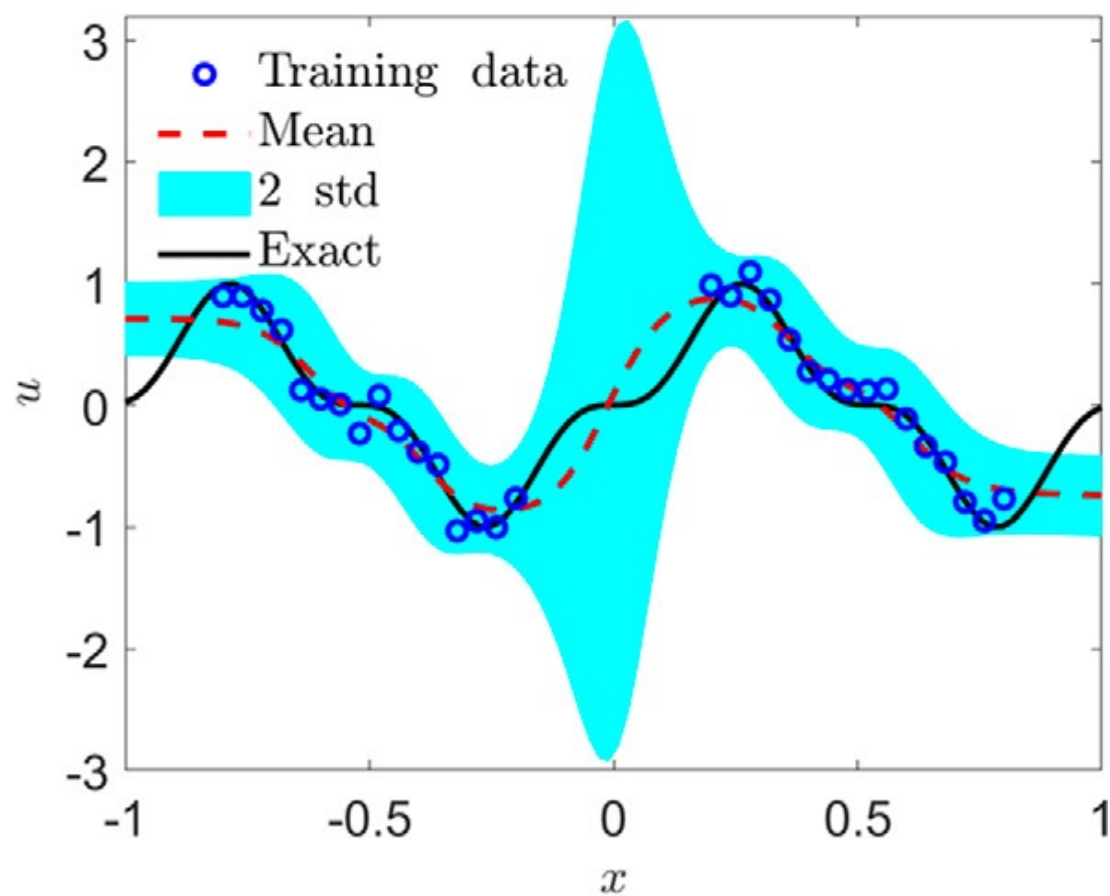
BNN-HMC



BNN-HMC

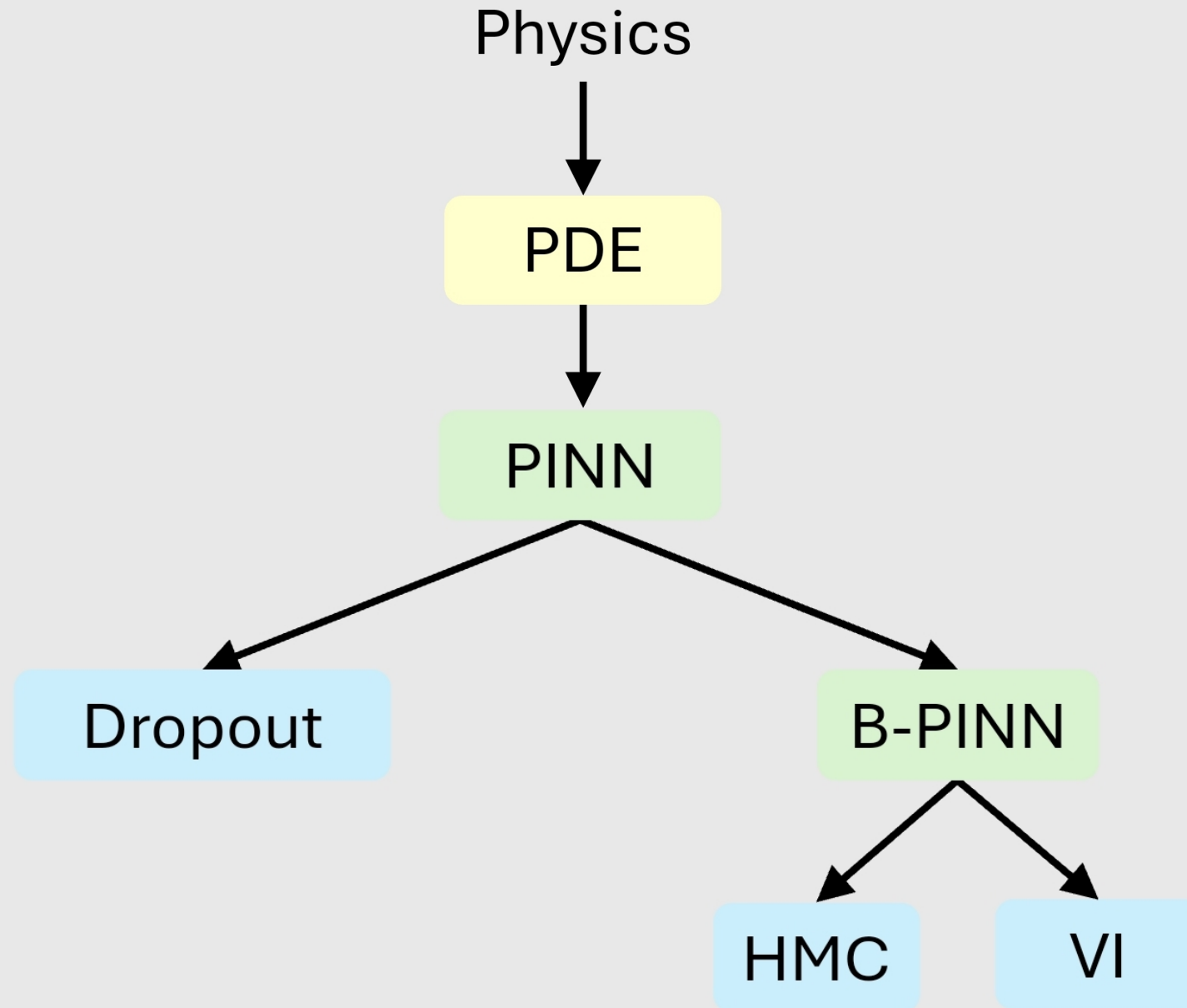


BNN-VI



From here on, B-PINNs
instead of BNNs are used





1D linear Poisson Equation

$$\mathcal{N}_x(\mathbf{u}; \lambda) = f, \quad x \in D$$

$$\mathcal{B}_x(\mathbf{u}; \lambda) = b, \quad x \in \Gamma$$



$$\lambda \partial_x^2 \mathbf{u} = f \quad x \in [-0.7, 0.7]$$

$$\mathbf{u} = b \quad x \in \{-0.7, 0.7\}$$

1D linear Poisson Equation

differential operator

potential field

charge density

$$\lambda \partial_x^2 u = f \quad x \in [-0.7, 0.7]$$
$$u = b \quad x \in \{-0.7, 0.7\}$$

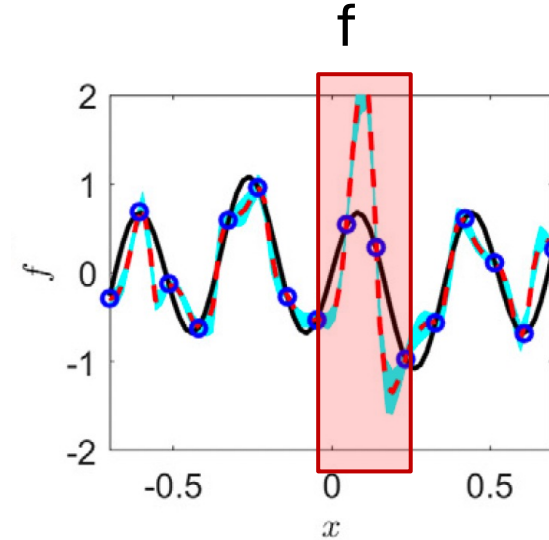
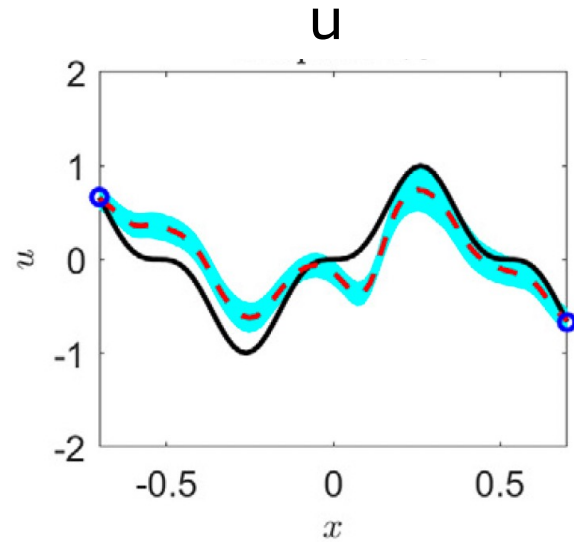
constant factor

$$u(x) = \sin^3(6x)$$

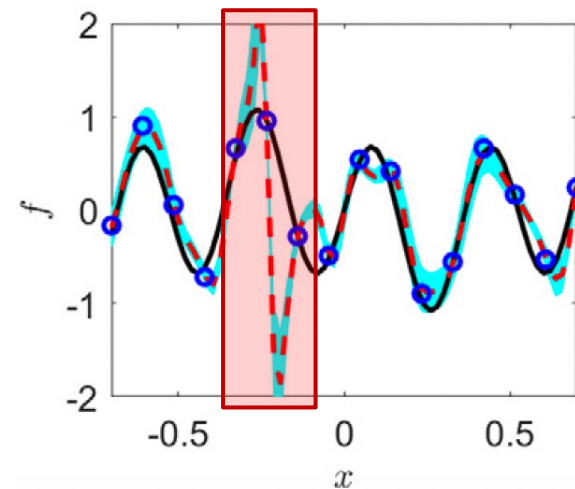
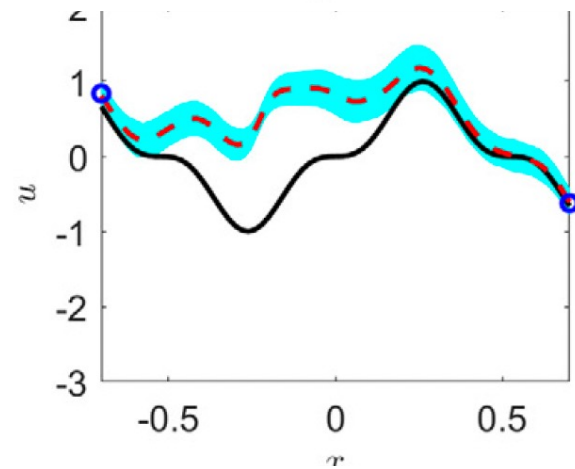
PINN-Dropout 1%

$$\lambda \partial_x^2 u = f$$

$$\epsilon_f, \epsilon_b \in \mathcal{N}(0, 0.01^2)$$



$$\epsilon_f, \epsilon_b \in \mathcal{N}(0, 0.1^2)$$

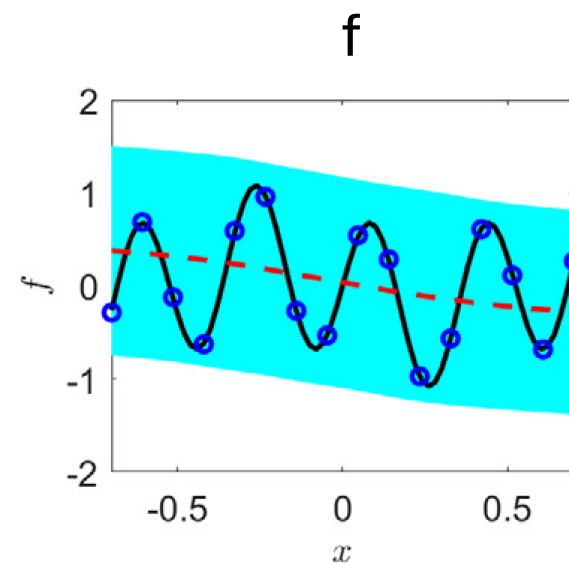
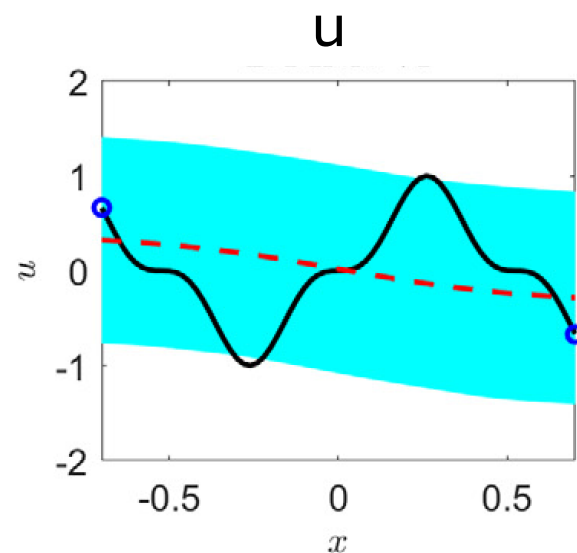


- Training data
- Mean
- 2 std
- Exact

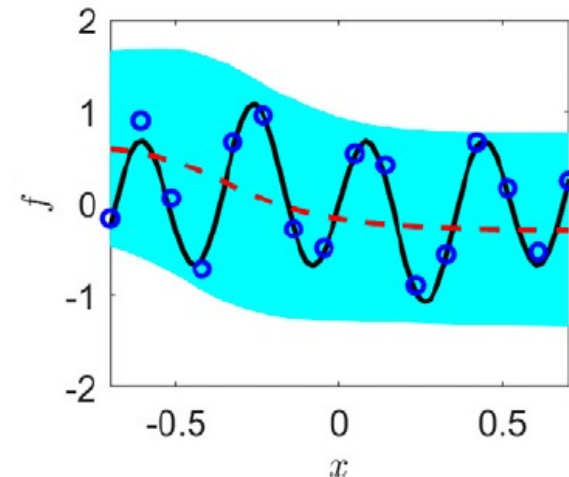
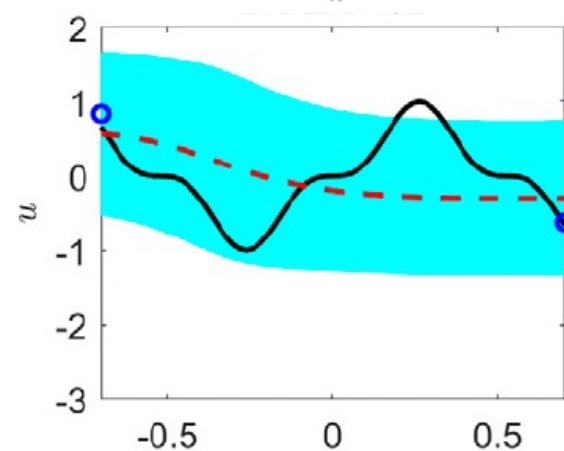
B-PINN-VI

$$\lambda \partial_x^2 u = f$$

$$\epsilon_f, \epsilon_b \in \mathcal{N}(0, 0.01^2)$$



$$\epsilon_f, \epsilon_b \in \mathcal{N}(0, 0.1^2)$$

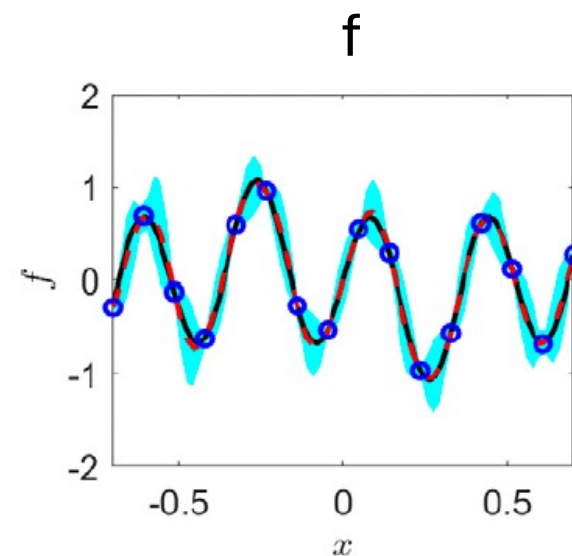
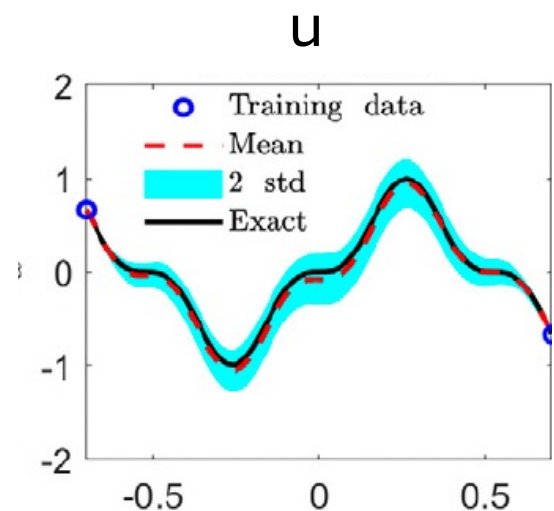


- Training data
- - - Mean
- 2 std
- Exact

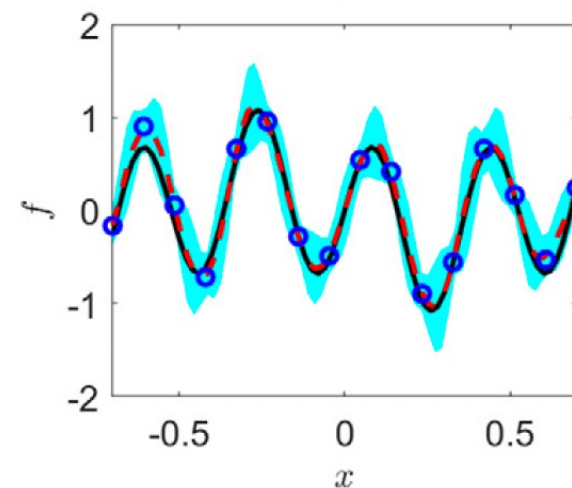
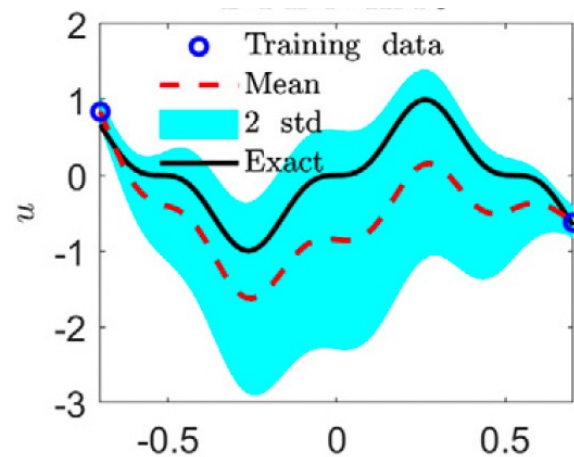
B-PINN-HMC

$$\lambda \partial_x^2 u = f$$

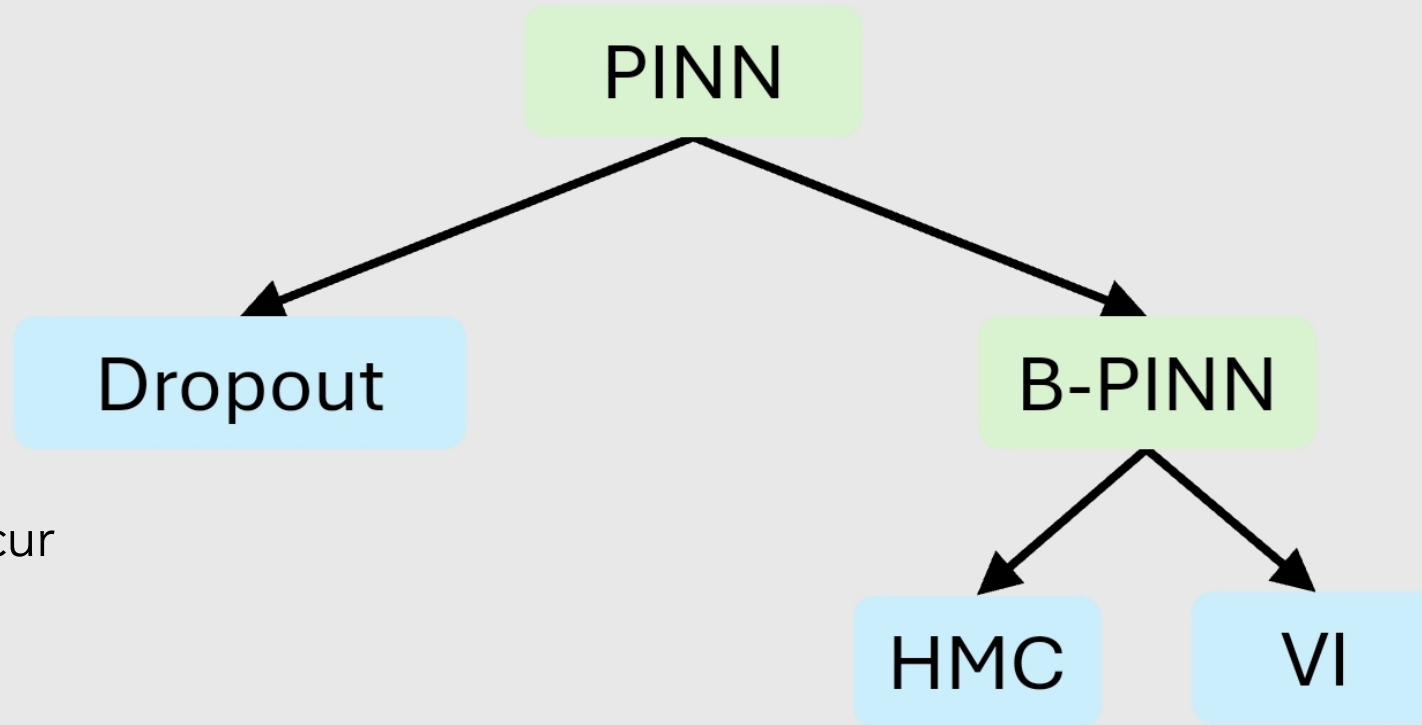
$$\epsilon_f, \epsilon_b \in \mathcal{N}(0, 0.01^2)$$



$$\epsilon_f, \epsilon_b \in \mathcal{N}(0, 0.1^2)$$



- Training data
- Mean
- 2 std
- Exact



- overfitting can occur

- accurate and robust,
even for noisy data

- failing can occur

1D nonlinear Poisson Equation

$$\lambda \partial_x^2 u = f \quad x \in [-0.7, 0.7]$$

$$u = b \quad x \in \{-0.7, 0.7\}$$



$$\lambda \partial_x^2 u + \underline{k \tanh(u)} = f \quad x \in [-0.7, 0.7]$$

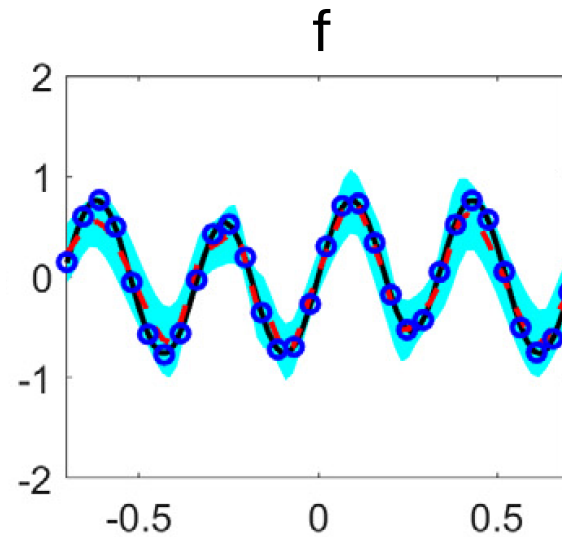
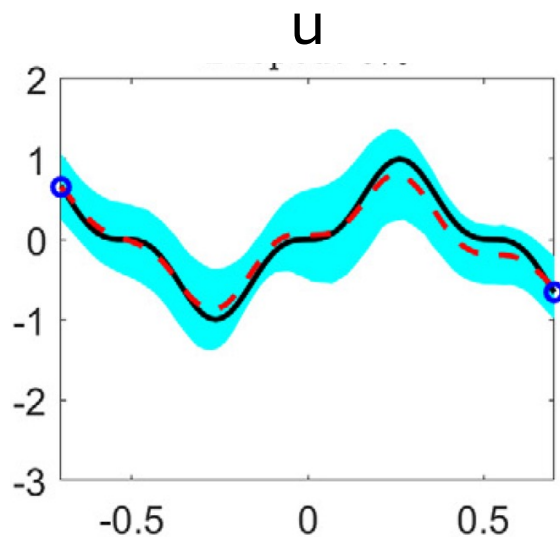
$$u = b \quad x \in \{-0.7, 0.7\}$$



PINN-Dropout 5%

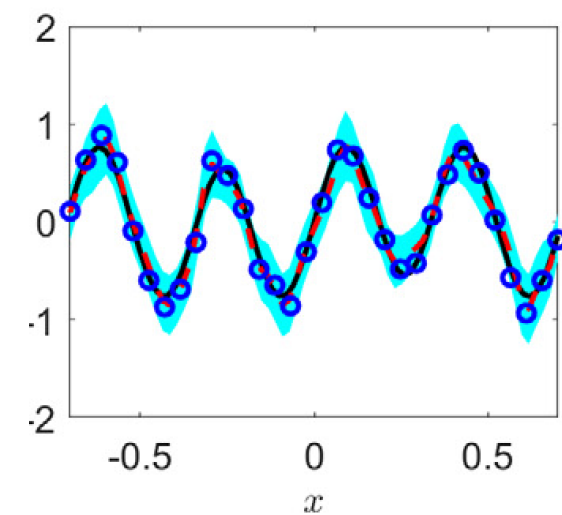
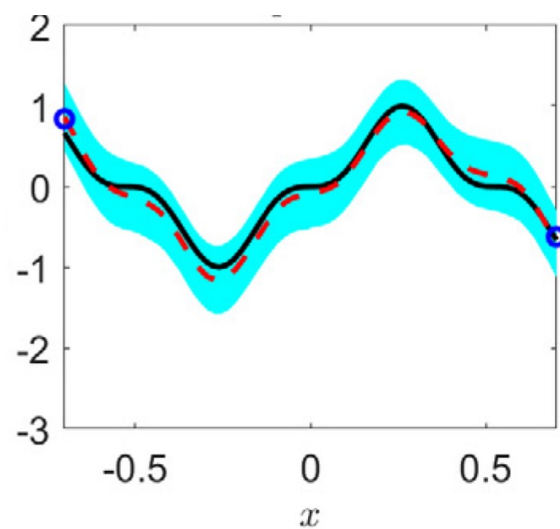
$$\lambda \partial_x^2 u + k \tanh(u) = f$$

$\epsilon_f, \epsilon_b \in \mathcal{N}(0, 0.01^2)$



- Training data
- - - Mean
- 2 std
- Exact

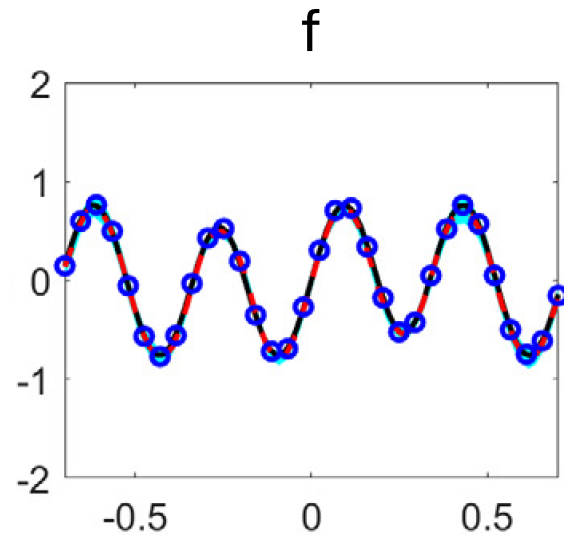
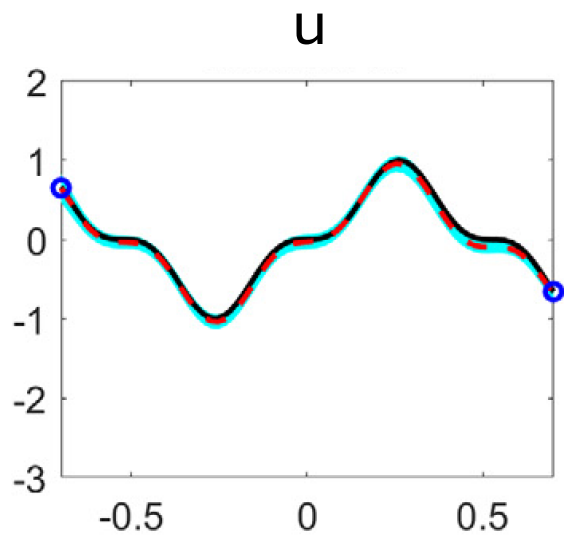
$\epsilon_f, \epsilon_b \in \mathcal{N}(0, 0.1^2)$



B-PINN-VI

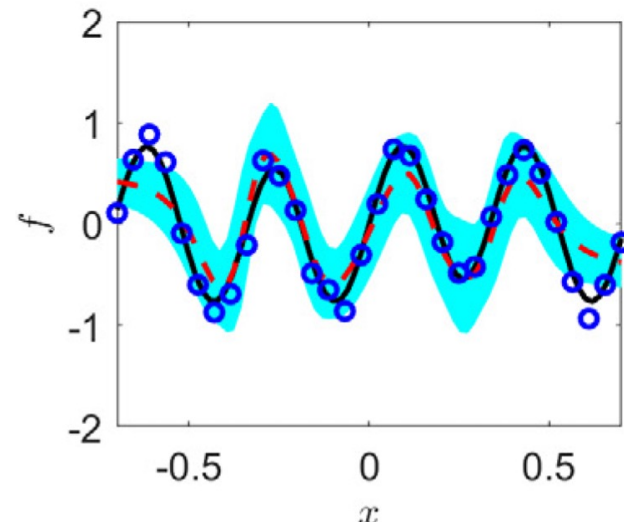
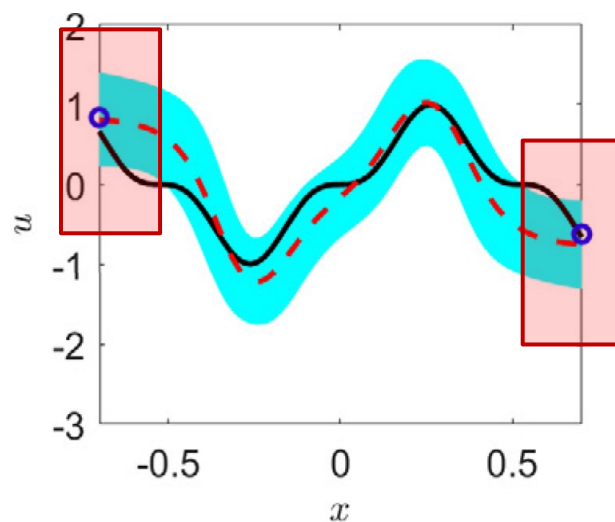
$$\lambda \partial_x^2 u + k \tanh(u) = f$$

$$\epsilon_f, \epsilon_b \in \mathcal{N}(0, 0.01^2)$$



- Training data
- - Mean
- 2 std
- Exact

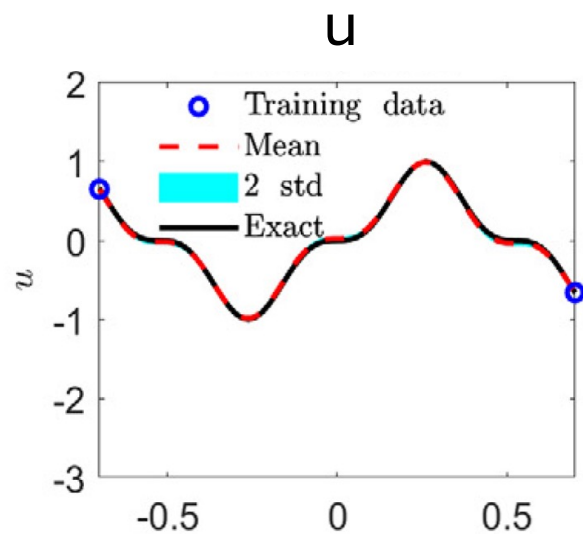
$$\epsilon_f, \epsilon_b \in \mathcal{N}(0, 0.1^2)$$



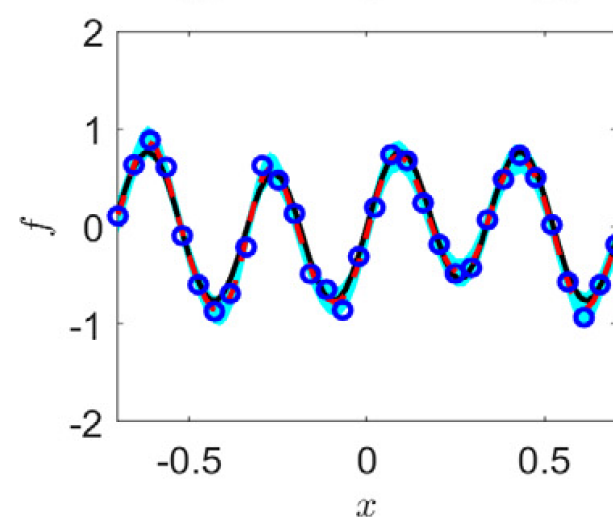
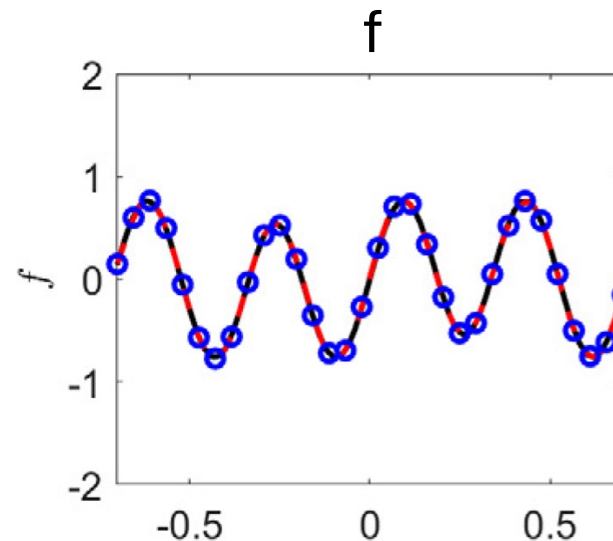
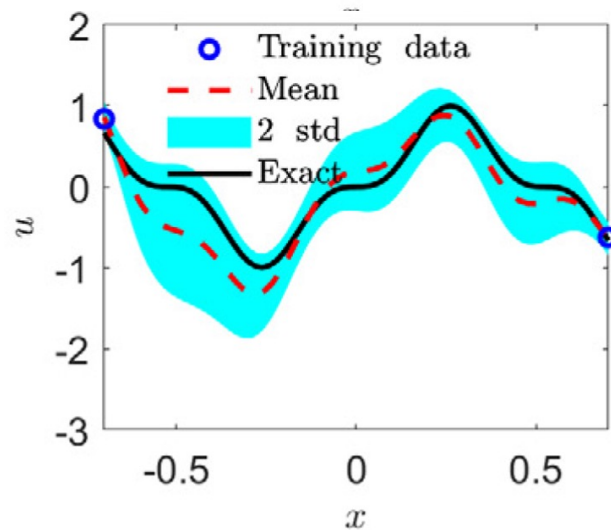
B-PINN-HMC

$$\lambda \partial_x^2 u + k \tanh(u) = f$$

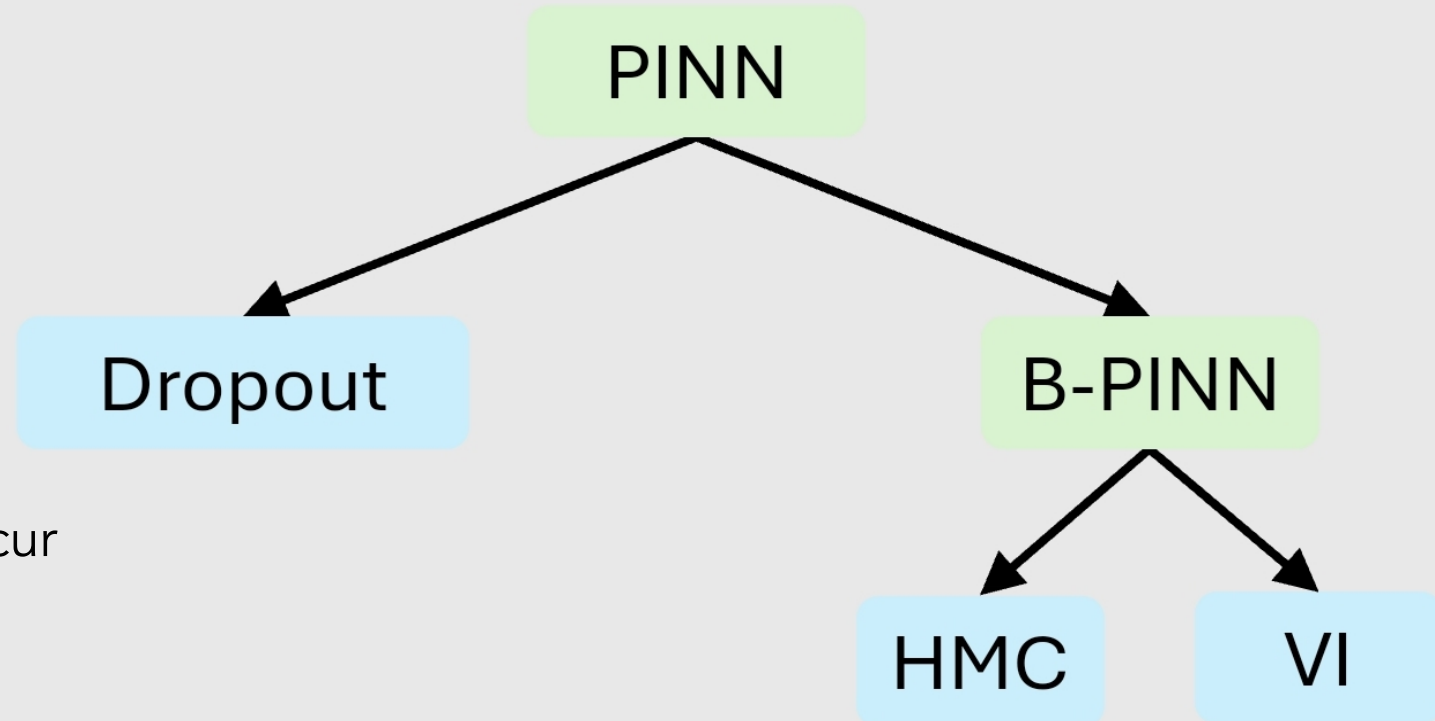
$$\epsilon_f, \epsilon_b \in \mathcal{N}(0, 0.01^2)$$



$$\epsilon_f, \epsilon_b \in \mathcal{N}(0, 0.1^2)$$

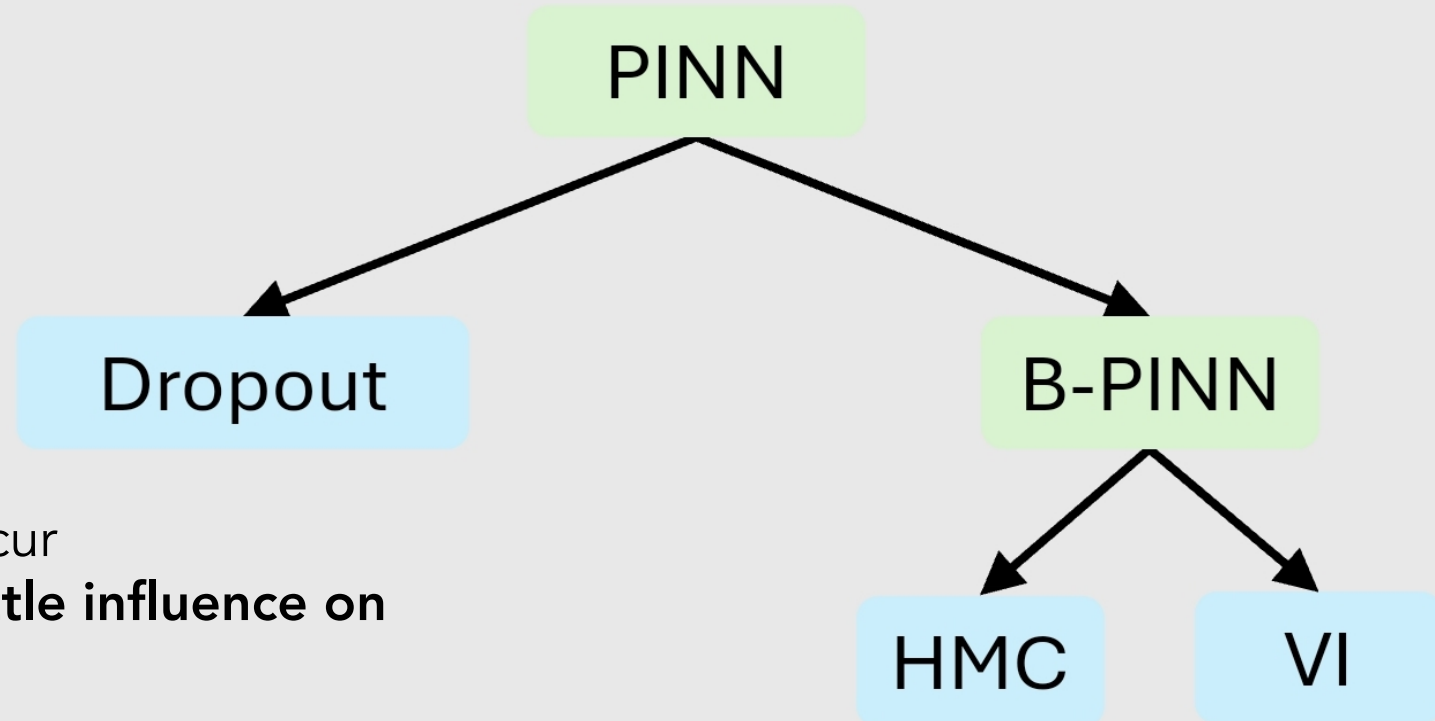


- Training data
- Mean
- 2 std
- Exact



- overfitting can occur

- accurate and robust, even for noisy data - failing can occur



- overfitting can occur
- **noise can have little influence on (un-)certainty**

- accurate and robust, even for noisy data

- failing can occur
- **unreasonable uncertainty at the boundaries**

1D nonlinear Poisson Equation - inverse problem

$$\begin{aligned} \lambda \partial_x^2 u + k \tanh(u) &= f & x \in [-0.7, 0.7] \\ u &= b & x \in \{-0.7, 0.7\} \end{aligned}$$



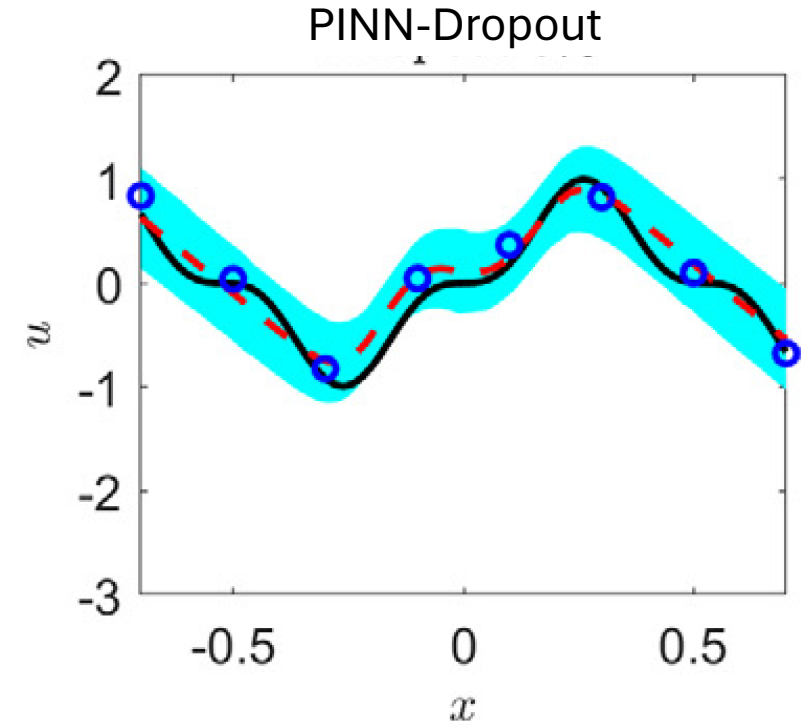
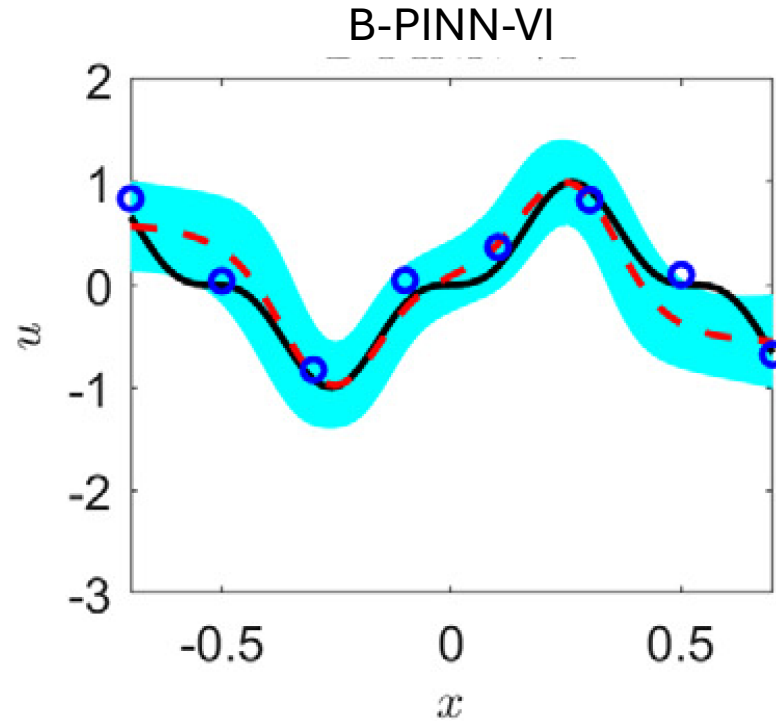
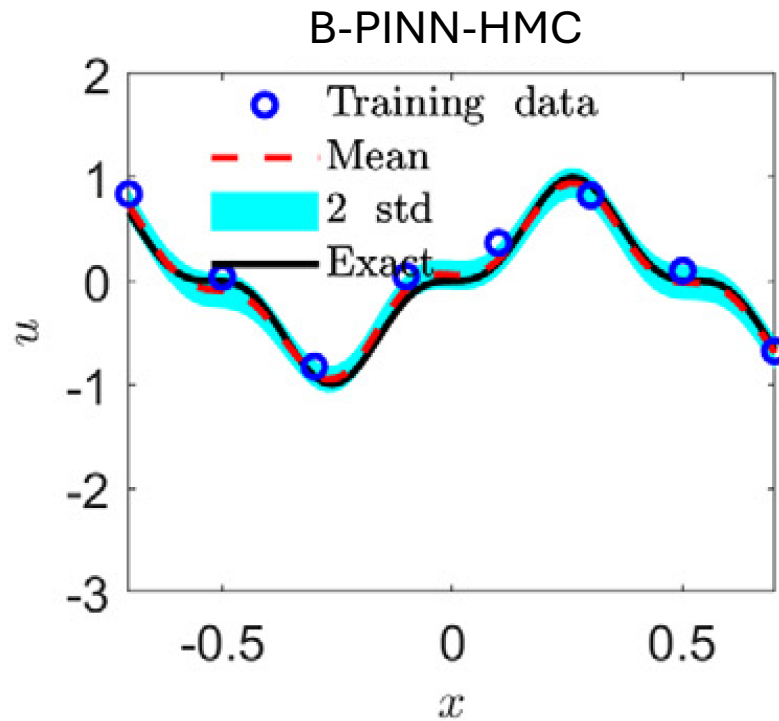
$$\begin{aligned} \lambda \partial_x^2 u + k \tanh(u) &= f & x \in [-0.7, 0.7] \\ u &= b & x \in \{-0.7, 0.7\} \end{aligned}$$



1D nonlinear Poisson Equation

$$\lambda \partial_x^2 u + k \tanh(u) = f$$

- inverse problem



$$\epsilon_f \sim \mathcal{N}(0, 0.1^2), \epsilon_u \sim \mathcal{N}(0, 0.1^2), \epsilon_b \sim \mathcal{N}(0, 0.1^2)$$

1D nonlinear Poisson Equation $\lambda \partial_x^2 u + k \tanh(u) = f$

- inverse problem

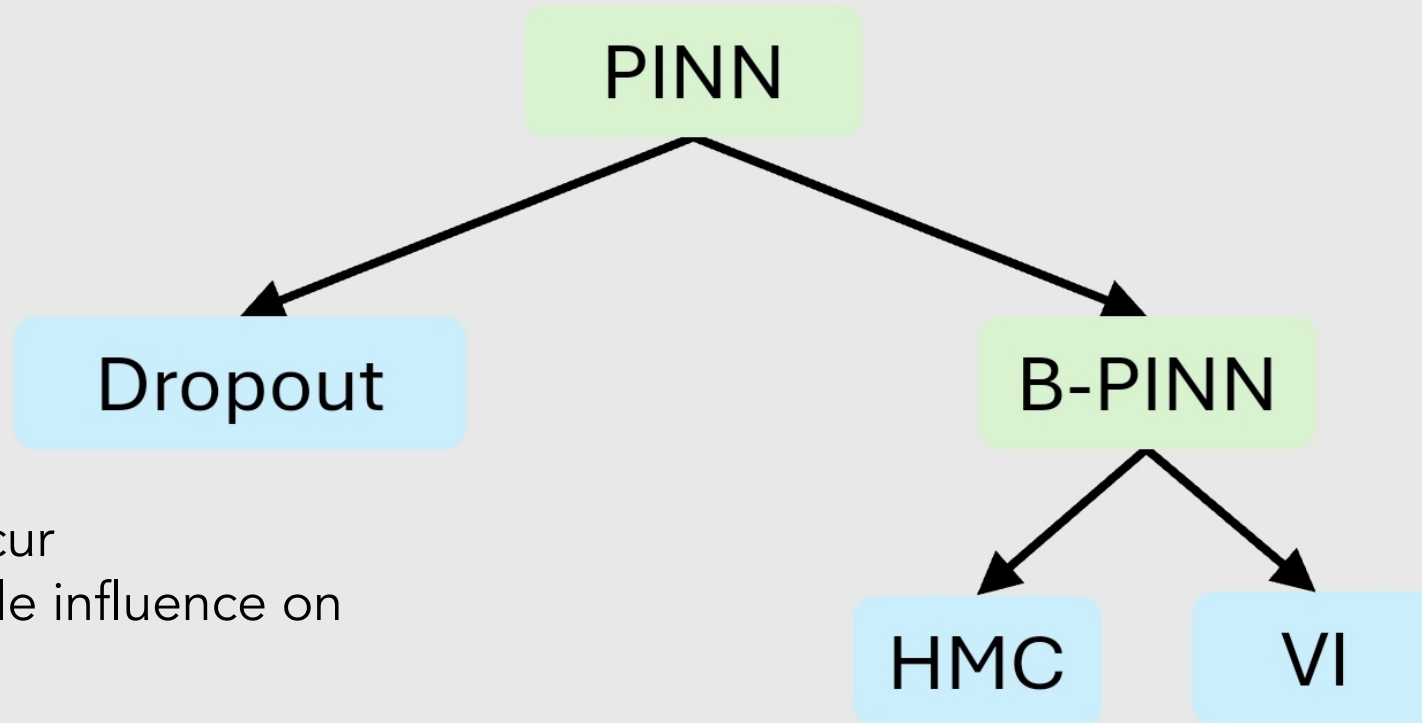
Exact value for k is 0.7

Noise scale		B-PINN-HMC	B-PINN-VI	Dropout-1%	Dropout-5%
0.01	Mean	0.705	0.708	0.714	0.669
	Std	5.75×10^{-3}	4.01×10^{-3}	4.38×10^{-3}	2.02×10^{-2}
0.1	Mean	0.665	0.775	0.746	0.633
	Std	5.63×10^{-2}	3.58×10^{-2}	6.508×10^{-3}	6.45×10^{-3}

quite accurate,
reasonable uncertainties

higher error than HMC,
reasonable uncertainties

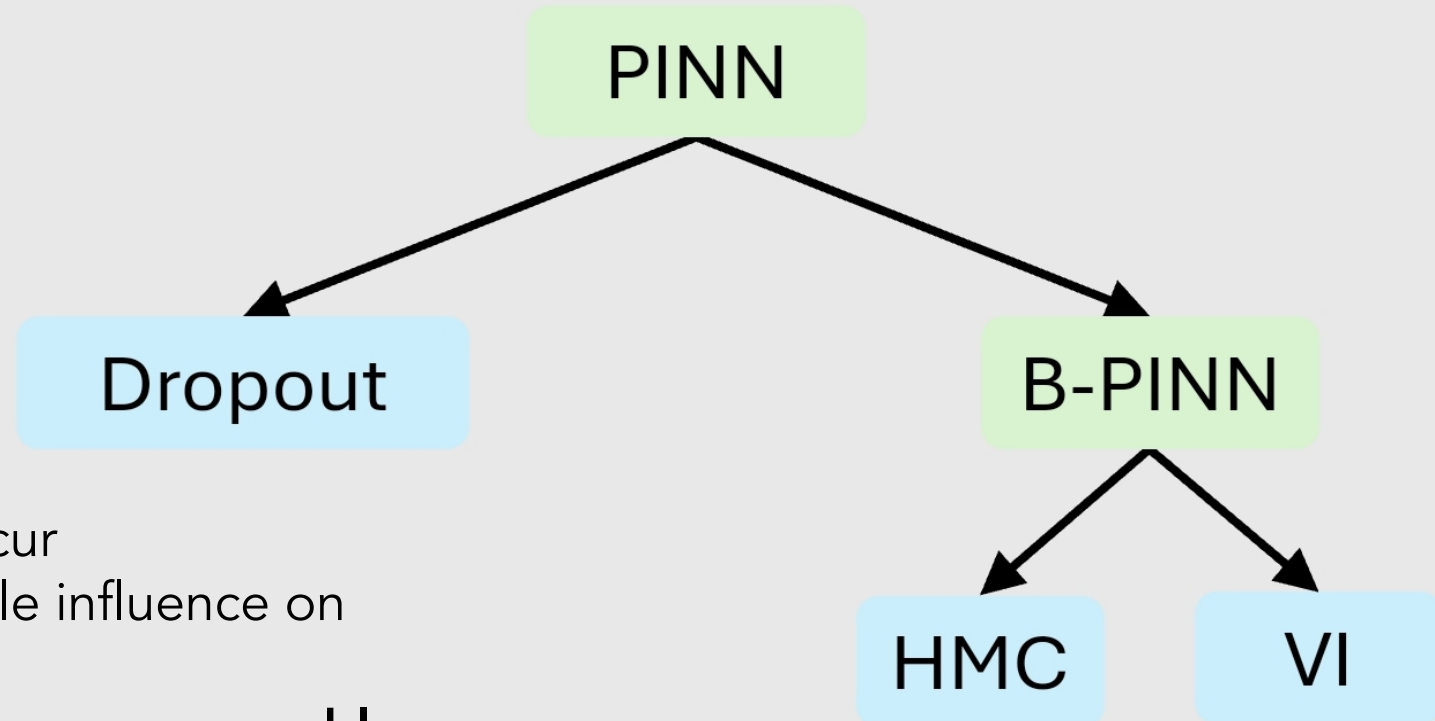
higher error than HMC and VI,
unreasonable uncertainties



- overfitting can occur
- noise can have little influence on (un-)certainty

- accurate and robust, even for noisy data

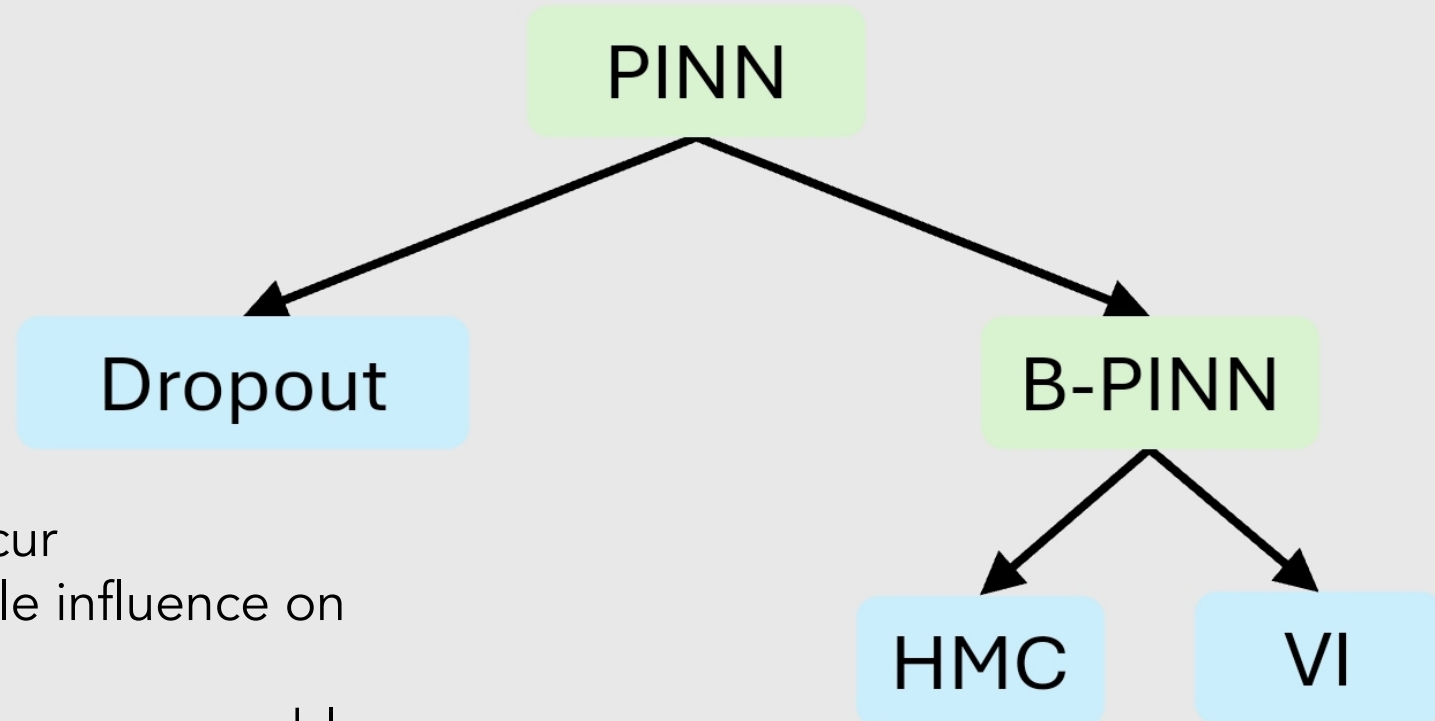
- failing can occur
- unreasonable uncertainty at the boundaries



- overfitting can occur
- noise can have little influence on (un-)certainty
- **noise can also have unreasonable influence on (un-)certainty of the model**

- accurate and robust, even for noisy data

- failing can occur
- unreasonable uncertainty at the boundaries
- **higher error than HMC**



- overfitting can occur
- noise can have little influence on (un-)certainty
- noise can also have unreasonable influence on (un-)certainty

- accurate and robust, even for noisy data

- failing can occur
- unreasonable uncertainty at the boundaries
- higher error than HMC

Open problems:

- Choice of e.g. prior distribution
- Big data case

EFFICIENT BAYESIAN PHYSICS INFORMED NEURAL NETWORKS FOR INVERSE PROBLEMS VIA ENSEMBLE KALMAN INVERSION

ANDREW PENSONEAULT* AND XUEYU ZHU†

Abstract. Bayesian Physics Informed Neural Networks (B-PINNs) have gained significant attention for inferring physical parameters and learning the forward solutions for problems based on partial differential equations. However, the overparameterized nature of neural networks poses a computational challenge for high-dimensional posterior inference. Existing inference approaches, such as particle-based or variance inference methods, are either computationally expensive for high-dimensional posterior inference or provide unsatisfactory uncertainty estimates. In this paper, we present a new efficient inference algorithm for B-PINNs that uses Ensemble Kalman Inversion (EKI) for high-dimensional inference tasks. By reframing the setup of B-PINNs as a traditional Bayesian inverse problem, we can take advantage of EKI's key features: (1) gradient-free, (2) computational complexity scales linearly with the dimension of the parameter spaces, and (3) rapid convergence with typically $\mathcal{O}(100)$ iterations. We demonstrate the applicability and performance of the proposed method through various types of numerical examples. We find that our proposed method can achieve inference results with informative uncertainty estimates comparable to Hamiltonian Monte Carlo (HMC)-based B-PINNs with a much reduced computational cost. These findings suggest that our proposed approach has great potential for uncertainty quantification in physics-informed machine learning for practical applications.

2.2. Hamiltonian Monte Carlo (HMC). Next, we briefly review Hamiltonian Monte Carlo (HMC), a popular inference algorithm for B-PINNs [48] that serves as a baseline method for our proposed method. Hamiltonian Monte Carlo (HMC)

[48] Liu Yang, Xuhui Meng, and George Em Karniadakis. B-PINNs: Bayesian physics-informed neural networks for forward and inverse PDE problems with noisy data. *Journal of Computational Physics*, 425:109913, jan 2021.